Unified Stochastic Geometry Model for MIMO Cellular Networks with Retransmissions

Laila Hesham Afify, Hesham ElSawy, Tareq Y. Al-Naffouri, and Mohamed-Slim Alouini

King Abdullah University of Science and Technology (KAUST)
Email: {laila.afify, hesham.elsawy, tareq.alnaffouri, slim.alouini}@kaust.edu.sa

Abstract—This paper presents a unified mathematical paradigm, based on stochastic geometry, for downlink cellular networks with multiple-input-multiple-output (MIMO) base stations (BSs). The developed paradigm accounts for signal retransmission upon decoding errors, in which the temporal correlation among the signal-to-interference-plus-noise-ratio (SINR) of the original and retransmitted signals is captured. In addition to modeling the effect of retransmission on the network performance, the developed mathematical model presents twofold analysis unification for MIMO cellular networks literature. First, it integrates the tangible decoding error probability and the abstracted (i.e., modulation scheme and receiver type agnostic) outage probability analysis, which are largely disjoint in the literature. Second, it unifies the analysis for different MIMO configurations. The unified MIMO analysis is achieved by abstracting unnecessary information conveyed within the interfering signals by Gaussian signaling approximation along with an equivalent SISO representation for the per-data stream SINR in MIMO cellular networks. We show that the proposed unification simplifies the analysis without sacrificing the model accuracy. To this end, we discuss the diversity-multiplexing tradeoff imposed by different MIMO schemes and shed light on the diversity loss due to the temporal correlation among the SINRs of the original and retransmitted signals. Finally, several design insights are highlighted.

Keywords—MIMO cellular networks, error probability, outage probability, ergodic rate, stochastic geometry, network design.

I. INTRODUCTION

Multiple-input-multiple-output (MIMO) transmission offers diverse options for antenna configurations that can lead to different diversity and multiplexing tradeoffs, which can be exploited to improve several aspects in wireless networks performance. For instance, link capacity gains are harvested by multiplexing several data streams into the same channel via MIMO spatial multiplexing. Enhanced link reliability is obtained by transmit and/or receive diversity. The network capacity is improved by accommodating more users equipment (UEs) per channel via multi-user MIMO techniques. Last but not least, enhanced interference management is achieved via beamforming or interference alignment schemes that suppress dominant interference sources at the receivers side, which improves the signal-to-interference-plus-noise ratio (SINR).

Motivated by its potential gains, MIMO is considered an essential ingredient in modern cellular networks and 3GPP standards to cope with the ever-growing capacity demands. However, the MIMO operation is understood and its associated gains are quantified for elementary network settings [1]–[3], which do not directly generalize to cellular networks. The operation of large-scale cellular network is highly affected by inter-cell interference, which emerges from spatial frequency reuse. Therefore, to characterize MIMO operation and quantify its potential gains in cellular network, the impact of per-base station (BS) precoding on the aggregate interference as well as the effect of the aggregate interference on the received SINR after MIMO post-processing should be characterized. Exploiting recent advances in stochastic geometry analysis, several mathematical frameworks are developed to study MIMO operation in cellular networks [4]–[20]. Stochastic geometry does not only provide systematic and tractable framework to model MIMO operation in interference environments, it also captures the behavior of realistic cellular networks as reported in [21]–[25]. The authors in [4] study the SINR coverage probability of orthogonal space-time block codes (OSTBC). Studies for the outage probability and ergodic rate for space-division-multiple-access (SDMA) MIMO, also known as multi-user MIMO (MU-MIMO), are available in [14], [15] for single-tier cellular networks and in [20] for multi-tier cellular networks. Coverage probability improvement via maximum-ratio-combining (MRC) with spatial interference correlation is quantified in [5]. The potential gains of beamforming and interference alignment in terms of SINR coverage and network throughput are quantified in [6]–[8]. Coverage probability and rate for MRC and optimum combining in uplink MIMO cellular networks are studied in [19]. Network MIMO via BS cooperation performance in terms of outage probability and ergodic rate are studied in [9]–[13]. Average symbol error probability (ASEP) and average pairwise error probability (APEP) for several MIMO configurations are studied in [16], [17]. Asymptotic analysis for minimum mean square error MIMO receivers is conducted in [18]. A general framework that can compare outage probability and rate for different MIMO schemes in advanced cellular network models is developed in [25].

Despite that the mathematical models presented in [4]–[18] are all based on stochastic geometry, there are significant differences in terms of the analysis steps as well as the level of details provided by each model. The majority of the models focus on the outage probability and ergodic capacity for simplicity [4]–[15], [19], [20], [25]. While both outage probability and ergodic rate are fundamental key performance indicators (KPIs) in wireless communication, they convey no information about the underlying modulation scheme, constellation size, or receiver type. Considering more tangible KPIs, such as decoding error probability and average throughput, requires alternative and more involved analysis as shown in [16], [17]. The error probability analysis

---

This work was funded by a CRG3 grant ORS#221 from the Office of Competitive Research (OCRF) at King Abdullah University of Science and Technology (KAUST).

---

1Throughput is defined as the number of successfully transmitted bits per channel use.
This article has been accepted for publication in a future issue of this journal, but has not been fully edited. Content may change prior to final publication. Citation information: DOI 10.1109/TWC.2016.2616394, IEEE Transactions on Wireless Communications

The main contributions of the developed framework can be summarized in the following points:

1. Developing a unified mathematical model that bridges the gap between error probability, outage probability, and ergodic rate analysis. Hence, it is possible to look at all three performance metrics within a single study.

2. Providing simplified analysis and computational complexity reduction while maintaining the model accuracy, when compared to the exact analysis in [16].

3. Accounting for the SINR temporal correlation and quantifying the resulting diversity loss.

4. Revealing the multiplexing cost, in terms of outage probability and decoding error, in large-scale cellular networks. We also show the appropriate diversity compensation for such cost.

5. Proposing a reliable network design strategy that is capable of appropriately adjusting the network parameters to meet desired design criteria.

A. Organization & Notation

The paper is organized as follows. Section II presents the generic system model for a downlink cellular network deploying an arbitrary MIMO setup. In Section III the Gaussian signaling approximation and the equivalent SISO-SINR model are presented. The unified model with and without retransmission is presented in Section IV. Section V illustrates how to represent different MIMO schemes via the equivalent SISO-SINR model. The model validation, via Monte-Carlo simulations, and the key findings of the paper are presented in Section VI and the paper is concluded in Section VII. Throughout the paper, we use the following notations: small-case bold-face letters (x) denote column vectors, upper-case bold-face letters (X) denote matrices, (·)H and (·)T denote the transpose and Hermitian operators, respectively. || · || is the Euclidean norm operator. 

\[ \mathbb{E}_x[·] \text{ and } \text{Var}_x[·] \text{ are the expectation and variance computed with respect to (w.r.t) the random variable } x, \text{ respectively, and } \text{erfc}(x) = \frac{2}{\sqrt{\pi}} \int_x^\infty e^{-t^2} \, dt \text{ is the complementary error function.} \]

II. System Model

A. Network and Propagation Models

We consider a single-tier downlink3 cellular network, where the BS locations are modeled by a homogeneous Poisson point process (PPP) \( \Psi_B \) with intensity \( \lambda_B \). UEs are distributed according to an independent homogeneous PPP \( \Psi_u \) with intensity \( \lambda_u \). BSs and UEs are equipped by \( N_t \) and \( N_r \) collocated antennas, respectively. Conditions on the relation between \( N_t \) and \( N_r \) depend on the MIMO setup under study, as will be shown later. Without loss of generality, we assume that \( \Psi_B = \{ r_0, r_1, r_2, \ldots \} \) contains the ascending ordered distances of the BSs from the origin (i.e., \( r_0 < r_1 < r_2 \)) and that the analysis is conducted on a test user located at the origin [26], [35]. Assuming nearest BS association, the test user is subject to interference from the BSs in \( \Psi^o = \Psi_B \setminus r_0 \), in which the distance between the test user and its serving BS has the probability distribution

\[ P_{\text{distr}}(r) \]

\[ \text{erfc}(x) = \frac{2}{\sqrt{\pi}} \int_x^\infty e^{-t^2} \, dt \]

2This paper is an extension of the work reported in [28] and presented in IEEE International Conference on Communications (ICC), 2016.

3The proposed analysis can be applicable to the uplink transmission with power control, following [19], [29], [33], [34].

4Colocated antennas is a common assumption in MIMO models based on stochastic geometry analysis to maintain the model tractability [4]–[8], [14]–[20].
function (PDF) $f_{r_0}(r) = 2\pi \lambda_B r e^{-\pi \lambda_B r^2}$, $r_0 > 0$. Let $p$ be the independent transmission probability for each BS in $\Psi^o$, then, the joint process of the active interfering BSs $\Psi^o \subset \Psi^A$ after independent thinning is also a PPP but with intensity $\lambda = p\lambda_B$ [35]. This assumption is used to reflect load awareness and/or frequency reuse as discussed in [16], [36]. Note that $p$ can be calculated as in [37], and setting $p = 1$ gives the traditional saturation condition (i.e., $\lambda_u \gg \lambda$) where all BSs are active.

A distance-dependent power-law path-loss attenuation is employed, in which the signal power attenuates at the rate $r^{-\eta}$ with the distance $r$, where $\eta > 2$ is the path-loss exponent. In addition to path-loss attenuation, we consider a Rayleigh multi-path fading environment between transmitting and receiving antennas, such that fading channels are independent from each other. That is, the channel gain matrix from a transmitting BS to a generic UE, denoted as $\mathbf{H} \in \mathbb{C}^{N_r \times N_t}$, has independent zero-mean unit variance complex Gaussian entries, such that $\mathbf{H} \sim \mathcal{CN}(0, \mathbf{I})$, where $\mathbf{I}$ is the identity matrix.

### B. Downlink MIMO Received Signal Model

For a general MIMO setup in Rayleigh fading environment, with arbitrary precoding/combining schemes, the complex baseband received signal vector accounting for the precoding matrices is denoted as $\mathbf{y}$, and after applying combining matrices is expressed as $\mathbf{y'}$ where

$$\mathbf{y'} = \sqrt{\frac{P}{r_0}} \mathbf{w}_o \mathbf{H}_i \mathbf{V}_i s_i + \mathbf{w}_o \mathbf{n}, \quad (1)$$

where $P = \frac{E_s}{N_0}$ is the transmit power per antenna at the BSs such that $E_s$ is the energy per symbol, $\mathbf{H}_i \in \mathbb{C}^{N_r \times N_t}$ is the useful channel matrix from the serving BS, and $\mathbf{H}_i \in \mathbb{C}^{N_r \times N_t}$ is the interfering channel matrix from the $i$th interfering BS, $\mathbf{H}_o$, and $\mathbf{H}_i$, have independent and identically distributed (i.i.d.) $\mathcal{CN}(0, 1)$ entries, $s_i \in \mathbb{C}^{L \times 1}$ and $s_i \in \mathbb{C}^{L \times 1}$ are, respectively, the intended and interfering symbols vectors, where $L$ represents the number of multiplexed data streams. The symbols in $s$ and $s_i$ are independently drawn from an equiprobable two dimensional unit-energy constellations. The matrices $\mathbf{V}_o \in \mathbb{C}^{N_r \times L}$ and $\mathbf{V}_i \in \mathbb{C}^{N_r \times L}$ are the intended and interfering precoding matrices at the intended BS and the $i$th interfering BS, respectively, while $\mathbf{W}_o \in \mathbb{C}^{L \times N_r}$ is the combining matrix at the test receiver. Note that $\mathbf{W}_o$ and $\mathbf{V}_i$ are tailored to each realization of $\mathbf{H}_i$, and are determined based on the employed MIMO scheme at the serving BS. On the other hand, $\mathbf{V}_i$ depends on the employed MIMO scheme at the $i$th interferer BS and is independent from the interfering channel matrix $\mathbf{H}_i$, as we assume no inter-BS interference management. $\mathbf{n} \in \mathbb{C}^{N_r \times 1}$ is the zero-mean additive white Gaussian noise vector with covariance matrix $\mathbf{N}_r \mathbf{I}_{N_r}$, where $\mathbf{I}_{N_r}$ is the identity matrix of size $N_r$. Last, we assume a per-symbol maximum likelihood (ML) receiver at the test UE to decode the symbols in $\mathbf{y}$.

### III. GAUSSIAN SIGNALLING APPROXIMATION & EQUIVALENT SISO REPRESENTATION

Assuming per-stream symbol-by-symbol ML receiver, the employed precoding, combining, and equalization techniques decouple symbols belonging to different streams (i.e., in case of multiplexing) and/or combine symbols belonging to the same stream (i.e., in case of diversity) at the decoder to allow disjoint and independent symbol detection across the multiplexed data streams. Hence, the precoding and combining matrices ($\mathbf{W}_o$ and $\mathbf{V}_i$) are tailored to each realization of $\mathbf{H}_i$, such that the product $\mathbf{W}_o \mathbf{H}_i \mathbf{V}_i$ always gives a diagonal matrix of size $L$, with the appropriate, possibly different, diagonal values that correspond to the stream being decoded. Without loss of generality, let us focus on the decoding performance of a generic symbol in the $l$th stream, in which the instantaneous received signal after applying combining/precoding techniques is given by

$$y_l = \sqrt{\frac{P}{r_0}} \mathbf{w}_{o,l}^T \mathbf{H}_i \mathbf{V}_{i,l} s_l + \mathbf{w}_{o,l}^T \mathbf{n}, \quad (2)$$

such that $l \in \{1, \cdots, L\}$, $\mathbf{w}_{o,l}$ is the $l$th column of matrix $\mathbf{W}_o^T$, and $\mathbf{V}_{i,l}$ is the $k$th column of matrix $\mathbf{V}_i$ which is designed based on the channel matrix between the $i$th interfering BS and its associated users, denoted as $\mathbf{H}_i$. Further, $g_{o,l}$ is a real random scaling factor that appears in the intended signal due to the equalization applied to detect the $l$th desired symbol, $\tilde{a}_{l}^{(i)}$, $\forall k$ are the complex random coefficients combining the interfering symbols from the antennas of the $i$th interfering BS. As shown in (2), the coefficients $\tilde{a}_{l}^{(i)}$ are generated from the product $\mathbf{w}_{o,l}^T \mathbf{H}_i \mathbf{V}_{i,l}$, which capture per-BS precoding and test receiver combining effects on the aggregate interference. As discussed earlier, $\mathbf{W}_o$ and $\mathbf{V}_i$ are designed independently from each other and neither of them accounts for the realization of the interfering channel matrix $\mathbf{H}_i$. Therefore, the coefficients $\tilde{a}_{l}^{(i)}$ randomly change with each realization of $\mathbf{H}_o$, $\mathbf{H}_i$, and $\mathbf{H}_i$.

It is clear, from (2), that the aggregate interference seen at the decoder of the test UE is affected by the per-interfering BS precoding scheme ($\mathbf{V}_{i,l}$), the number of streams transmitted by each BS ($L$), the per-stream transmitted symbol ($s_l$), and the employed combining technique ($\mathbf{w}_{o,l}$) at the test UE. Therefore, characterizing the aggregate interference in (2) is essential to characterize the MIMO operation in cellular networks. The aggregate interference term contains three main sources of randomness, namely, the network geometry, the channel gains, and the interfering symbols. The average decoding error performance is only characterized for certain distributions of additive noise channels (e.g., Gaussian [38], Laplacian [39], [40], and Generalized Gaussian [41]), in which the additive white Gaussian noise (AWGN) channel represents the simplest case. Accounting for the exact distribution of the interfering symbols, the interference-plus-noise distribution does not directly fit into any of the distributions where the average decoding error performance is known. Hence, averaging over the transmitted symbols cannot be directly conducted unless the interference-plus-noise term is expressed or approximated via one of the distributions where the average decoding error performance is known. The authors in [16] proposed the equivalent-in-distribution (EID)

---

5According to the employed MIMO setup, we might need to introduce a slight abuse of notation to preserve the convention used in (1). For instance, in multi-user MIMO setup with $K$ single-antenna users, the parameter $N_r$ should be replaced with $K$ so that the signal model in (1) remains valid.

6The precoding and combining matrices are functions of the channel gains.
approach where an exact conditional Gaussian representation for the aggregate interference in (2) is achieved. Hence, the conditional error probability analysis, conditioned on network geometry and channel gains, is conducted via error probability expressions for AWGN channels, followed by a deconditioning step. The main drawback of the EiD approach is that it requires characterizing the interference signals at the baseband level to achieve the conditional Gaussian representation, which complicated both averaging steps, especially in MIMO networks. Furthermore, the EiD approach for the error probability analysis in [16] is disjoint from the outage probability and ergodic rate analysis in [4], [6]–[12], [14], [15], [19], [20].

To facilitate the error probability analysis and achieve a unified error probability, outage probability and ergodic rate analysis, we only account for the entries in the intended symbol vector \( \mathbf{s} \) of (1) and abstract the entries in \( \mathbf{s} \), by i.i.d. zero-mean Gaussian signals \( \tilde{\mathbf{s}} \), with unit-variance. Such abstraction ignores the unnecessary and usually unavailable combining/equalization can be represented via the following mean channel power gains and Gaussian signaling approximation, which complicated both averaging steps, especially in MIMO networks.

\[
\mathbf{y}_i = \sqrt{\frac{P}{r_i}} g_{o, i} s_i + \sum_{r_i \in \Phi} \sqrt{\frac{P}{r_i}} \sum_{k=1}^{L} a_{l, k}^{(i)} \tilde{s}_{i, k} + \mathbf{w}_{o, i} \mathbf{n}, \tag{3}
\]

It is important to note that the random variables \( g_{o, i} \) are i.i.d. and hence, there is no loss in generality to drop the index \( i \) and study an arbitrary stream. Conditioned on \( r_i \) and \( a_{l, k}^{(i)} \), the lumped interference-plus-noise term in (3) is Gaussian because of the Gaussianity of \( \tilde{s}_{i, k}, \forall i, k \). This renders the well-known AWGN error probability expressions legitimate to conduct the averaging over the transmitted symbols, in which the noise variance used in the AWGN-based expressions is replaced by the variance of the lumped interference and noise terms in (3). The Gaussian signaling approximation leads to the following proposition.

**Proposition 1:** Consider a downlink MIMO cellular network with \( N_t \) antennas at each BS and \( N_r \) antennas at each UE in a Rayleigh fading environment with i.i.d. unit-mean channel power gains and Gaussian signaling approximation for the interfering symbols, then the per-data stream conditional SINR at the decoder of a generic UE after combining/equalization can be represented via the following equivalent SISO-SINR

\[
\Upsilon = \frac{P r_i^{-\eta} g_{o, i}^2}{\sum_{r_i \in \Phi} P r_i^{-\eta} g_{i} + N_o}, \tag{4}
\]

where the random variables \( g_{o, i} \sim \text{Gamma}(m_o, 1) \) and \( g_{i} \sim \text{Gamma}(m_i, 1) \) capture the effect of MIMO precoding, combining, and equalization. The values of \( m_o \) and \( m_i \) are determined based on the number of antennas (\( N_t \) and \( N_r \)), the number of multiplexed data streams per BS (\( L \)) and the employed MIMO configuration as shown in Table I.

**Proof:** The detailed discussion and proof of each MIMO setup is given in Section V in the corresponding lemma shown in Table I. Here, we just sketch a high-level proof of the proposition. The equivalent channel gains in (4) are \( \tilde{g}_o = |g_{o, i}|^2 \) and \( \tilde{g}_i = |g_{i}|^2 = \sum_{k=1}^{L} |a_{l, k}^{(i)}|^2 \), where \( g_{o, i} \) is the random scale for the intended symbol due to precoding and combining/equalization as shown in (2). Since preceding and combining/equalization are usually in the form of linear combination of the channel power gains and that the channel gains have independent Gaussian distributions, both random variables \( \tilde{g}_o \) and \( \tilde{g}_i \) are independent \( \chi^2 \)-distributed with degrees of freedom equal to the number of linearly combined random variables, which depends on the number of antennas, precoding technique, and number of multiplexed data streams per BS. Note that the \( \chi^2 \) distribution for interfering channel gains \( \tilde{g}_i \) is exact only if the precoding vectors in each BS are independent. In the case of dependent precoding vectors, the correlation is ignored and the \( \chi^2 \) distribution for \( \tilde{g}_i \) is an approximation. Such approximation is commonly used in the literature for tractability [4], [6]–[11], [14]–[17], and is verified in Section VI. Exploiting the one-to-one mapping between the \( \chi^2 \) distribution and the gamma distribution, we follow the convention in [4], [6]–[11], [14]–[17] and use the gamma distribution, instead of the \( \chi^2 \) distribution, for \( \tilde{g}_o \) and \( \tilde{g}_i \).

It is important to note that, the proposed SISO-SINR model relies on the assumption of independent Rayleigh fading channels. While spatially correlated fading channels model provides more realistic fading environment [42], however, such correlation is ignored for tractability as in [4], [6]–[11], [14]–[17]. Proposition 1 gives the equivalent SISO representation for the MIMO cellular network in which the effect of precoding, combining, and equalization of the employed MIMO scheme is abstracted by the random variables \( g_{o, i} \) and \( g_{i} \), \( \forall i \) in (4). Hence, unified analysis and expressions for different KPIs and MIMO configurations, respectively, are viable as shown in the next section.

**IV. UNIFIED PERFORMANCE ANALYSIS**

Based on the Gaussian signaling approximation, interference-plus-noise in (3) is conditionally Gaussian. Hence, the decoding error performance of the MIMO scheme is studied by plugging the conditional SINR in (4) with the appropriate channel gains (i.e., \( g_{o, i} \) and \( g_{i} \)) in the corresponding AWGN-based decoding error expression, followed by an averaging over the channel gains and network geometry. Using the AWGN expression for the SEP for square quadrature amplitude modulation (\( M \)-QAM) scheme given in [38], the ASEPs in MIMO cellular networks can be expressed as

\[
\text{ASEP} (\Upsilon) = w_1 E \left[ \text{erfc} \left( \sqrt{\beta \Upsilon} \right) \right] + w_2 E \left[ \text{erfc}^2 \left( \sqrt{\beta \Upsilon} \right) \right], \tag{5}
\]

where \( w_1 = 2\sqrt{M-1}/\sqrt{M} \), \( w_2 = -\left( \sqrt{M-1}/\sqrt{M} \right)^2 \), and \( \beta = \frac{3}{2(M-1)} \) are constellation-size specific constants. The Gaussian signaling approximation is also the key that unifies the ASEP, outage probability, and ergodic rate analysis. This is because both outage and capacity are information theoretic KPIs that implicitly assume Gaussian codebooks, which directly lead to the conditional SINR in the form given by Proposition 1. Consequently, both the outage probability and ergodic capacity are also functions of the SINR in the form of (4), and are given by

\[
O = P \{ \Upsilon < \theta \}, \tag{6}
\]
TABLE I: SISO-equivalent gamma distribution parameters for various MIMO settings.

<table>
<thead>
<tr>
<th>MIMO Setup</th>
<th>L</th>
<th>m₁</th>
<th>m₂</th>
<th>Accuracy</th>
<th>Proof</th>
</tr>
</thead>
<tbody>
<tr>
<td>SISO</td>
<td>1</td>
<td>N₁</td>
<td>N₁</td>
<td>Exact</td>
<td>Lemma 3</td>
</tr>
<tr>
<td>OSTBC</td>
<td>N₁</td>
<td>N₁</td>
<td>N₁</td>
<td>Exact</td>
<td>Lemma 5</td>
</tr>
<tr>
<td>ZF-RS</td>
<td>N₁</td>
<td>N₁</td>
<td>N₁+1</td>
<td>Exact</td>
<td>Lemma 6</td>
</tr>
<tr>
<td>SDMA</td>
<td>K</td>
<td>N₁</td>
<td>K</td>
<td>Approx.</td>
<td>Lemma 7</td>
</tr>
<tr>
<td>MISO</td>
<td>1</td>
<td>N₁</td>
<td>1</td>
<td>Exact</td>
<td>Corollary 1</td>
</tr>
<tr>
<td>SM-MIMO</td>
<td>N₁</td>
<td>N₁</td>
<td>N₁</td>
<td>Approx.</td>
<td>Lemma 2</td>
</tr>
</tbody>
</table>

In order to derive the actual performance metrics from the conditional SINR expression in Section III, the expectations in (5), (6), and (7) are w.r.t the network geometry and channel gains, which are evaluated via stochastic geometry analysis. Such expectations are usually expressed in terms of the Laplace transform (LT)\(^8\) of the aggregate interference power in (4), denoted as \(\mathcal{L} = \sum_{r_i \in \Phi_\theta} P_{r_i}^{-\eta} g_i\). The LT of the interference power in the SISO-equivalent SINR given in (4) is characterized by the following lemma.

**Lemma 1:** Consider a cellular network with MIMO transmission scheme that can be represented via the equivalent SISO-SINR in (4) and BSs modeled via a PPP with intensity \(\lambda\), in which each BS transmits a symbol vector of length \(L\) per channel use (pcu) with symbols drawn from a zero-mean unit-variance Gaussian distribution, then the LT of the interference power affecting an arbitrary symbol at a receiver located \(r_0\) meters away from its serving BS is given by

\[
\mathcal{L}_{I|r_0}(z) = \exp\left\{-\pi \lambda r_0^2 \left[1 + \sum_{r_i \in \Phi_\theta} \frac{1}{z r_i^\eta} - 1\right]\right\},
\]  

(8)

where \(2F_1(\cdot; \cdot; \cdot)\) is the Gauss hypergeometric function [43].

**Proof:** Starting from the definition of the LT, we have

\[
\mathcal{L}_{I|r_0}(z) \overset{(a)}{=} \mathbb{E}\left[\exp\left\{-\sum_{r_i \in \Phi_\theta} P_{r_i}^{-\eta} g_i\right\}\right] = \exp\left\{-2 \pi \lambda \int_{r_0}^\infty \mathbb{E}\left[1 - e^{-z P x^{-\eta}}\right] x dz\right\},
\]

\[
\overset{(b)}{=} \exp\left\{-2 \pi \lambda \int_{r_0}^\infty \left[1 - \frac{1}{(1 + z P x^{-\eta})^{m_1}}\right] x dz\right\},
\]

(9)

where (a) follows from the probability generating functional (PGFL) of the homogeneous PPP [35] constituted by the interferers lying outside a disk of radius \(r_0\), i.e. \(r_i > r_0\), with intensity \(\lambda\), and (b) follows from the LT of the gamma distribution channel gains with shape parameter \(m_1\) and unity scale parameter. Solving the integral, yields the LT in (8). \(\blacksquare\)

Using Proposition 1 and Lemma 1, we arrive to the unified MIMO expressions for the ASEP, outage probability, and ergodic rate in the following theorem.

**Theorem 1:** Unified Analysis: Consider a cellular network with MIMO transmission scheme that can be represented via the equivalent SISO-SINR in (4) and BSs modeled via a PPP with intensity \(\lambda\), in which each BS transmits a symbol vector of length \(L\) pcu with symbols drawn from an equiprobable unit-power M-QAM modulation scheme, then the ASEP for an arbitrary symbol is approximated by (10), where \(1F_1(\cdot; \cdot; \cdot)\) is the Kummer confluent hypergeometric function [43].

For an interference-limited scenario, the probability that the average SIR (averaged over all symbols) for an arbitrary stream goes below a threshold \(\theta\) is given by

\[
O(\theta) = 1 - \int_0^\infty 2 \pi \lambda x e^{-x \lambda} B^2 \sum_{j=0}^{m_1-1} (-1)^j \frac{d^j}{dz^j} \left(\frac{\theta x^{\eta}}{P}\right) z^{m_1-1} dz, dx,
\]

(11)

and the ergodic rate for an arbitrary stream data rate is given by

\[
\mathcal{R} = \int_0^\infty \int_0^\infty 2 \pi \lambda x e^{-x \lambda} B^2 \mathcal{L}_{I|x}(z) \left(1 - \frac{1 + z}{z^{m_1}}\right) dz dx,
\]

(12)

where \(\mathcal{L}_{I|x}(z)\) in (10), (11), and (12) is the LT given in Lemma 1 when replacing \(r_0\) with \(x\).

**Proof:** See Appendix A.

The ASEP given in (10) is an approximation due to the Gaussian signaling abstraction used in (3). On the other hand, expressions for the outage probability and ergodic rate are exact, because both are typically derived in the literature based on the Gaussian codebooks assumption. The outage probability in (11) is given for interference-limited networks for tractability, which is a common assumption in cellular network because the interference term usually dominates the noise. Both equations (11) and (12) are approximations in cases of SDMA and SM-MIMO due to the approximate estimations of the interference as shown in Table I. It is meant to be mentioned that different exact/approximate forms for the outage probability and ergodic rate, shown in (11) and (12), respectively, can be derived via Gil-Pelaez inversion theorem as in [25], [44] and Alzer’s inequality as in [45], [46]. Nevertheless, both approaches are based on the Gaussian signaling approximation and equivalent SISO representation given in Proposition 1.

The ASEP expression given in (10) provides a unified ASEP expression for all considered MIMO schemes. Second, the ASEP is characterized based on the LT given in Lemma 1, which is the same LT used for characterizing the outage probability and ergodic rate. Third, the computational complexity to evaluate (10) is less than the complexity of the ASEP expressions in [16]. The reduced complexity of (10) is because it includes a single hypergeometric function in the exponential term while the expressions for the ASEP in [16] include summations of hypergeometric functions inside the exponential term. It should be highlighted that the EID approach leads to the same computational complexity as in (10) for some modulation schemes, e.g., phase-shift-keying (M-PSK) [16]. Nevertheless, the proposed Gaussian approximation and equivalent SISO representation simplifies the ASEP analysis and unifies it with outage and ergodic rate results when compared to the EID approach which always contains the complex baseband interference analysis.

A. The Effect of Temporal Correlation on Retransmissions

The network performance with retransmission cannot be directly deduced from Lemma 1 and Theorem 1 due to the temporal correlation interference. Despite that we assume that the channel fading independently changes from one time slot to another, the interference at a given location is correlated across time for the same network realization due to the fixed locations of the complete set of BSs. In other...
words, assuming a static UE, a subset of the interferers at a
given time slot might also be interfering in subsequent time
slots, which introduces temporal interference correlation that
needs to be taken into account. In this section, we study
temporal correlation between two arbitrary time slots at the
same spatial location. Based on the interference temporal
relation, we derive the conditional success probability of
a retransmission such that an earlier transmission at the same
location at two different time slots, denoted by
\( \lambda \).
Based on the aforementioned two steps, the equivalent
MIMO scheme at the first and second time slots, respectively,
and \( \mathbf{F}_1 \) is the Appell Hypergeometric function, which
extends the hypergeometric function to two variables
\( x \) and \( y \) [47].

\[ \tilde{\mathbf{Y}}_1 \) and \( \tilde{\mathbf{Y}}_2 \) are the SIRs at the first and second
transmissions. Using the joint LT in Lemma 2, the average
coverage probability with retransmission in MIMO cellular
network is given by the following theorem.

**Theorem 2:** Consider a cellular network with MIMO
transmission scheme that can be represented via the equiv-
elent SISO-SINR in (4) and BSs modeled via a PPP with
intensity \( \lambda \), the SIR coverage probability for a generic UE
with retransmission such that the serving BS and interfering
BSs may use different MIMO schemes across time, is given by
(15),

where \( \tilde{\theta} = \frac{\eta x y}{\pi} \).

\[ \tilde{\mathbf{Y}}_1 \) and \( \tilde{\mathbf{Y}}_2 \) are the SIRs at the first and second
transmissions. Using the joint LT in Lemma 2, the average
coverage probability with retransmission in MIMO cellular
network is given by the following theorem.

**Theorem 2:** Consider a cellular network with MIMO
transmission scheme that can be represented via the equiv-
elent SISO-SINR in (4) and BSs modeled via a PPP with
intensity \( \lambda \), the SIR coverage probability for a generic UE
with retransmission such that the serving BS and interfering
BSs may use different MIMO schemes across time, is given by
(15),

where \( \tilde{\theta} = \frac{\eta x y}{\pi} \).

Before giving numerical results and insights obtained from
the developed mathematical model, we first illustrate how the
equivalent SISO-SINR model given in Proposition 1 holds for
the considered MIMO schemes.

V. CHARACTERIZING MIMO CONFIGURATIONS

This section details the methodology to abstract different
MIMO configuration via the equivalent SISO model given in
Proposition 1 with parameters given in Table I. In order to
conduct the analysis for the different MIMO setups, we
first need to define the set \( \{ \tilde{\mathbf{H}} \} \) as the set of channel matrices
that affect the aggregate interference signals due to precoding
and/or combining. For instance, due to precoding, combining,
and equalization, the interference from the \( i \)th interfering BS
is multiplied by \( \mathbf{W}_i \mathbf{H}_i \mathbf{V}_i \), and hence, \( \{ \tilde{\mathbf{H}} \} = \{ \mathbf{H}_o, \mathbf{H}_i \} \),
where \( \mathbf{H}_o \) and \( \mathbf{H}_i \) are the channel matrices between, re-
spectively, the intended BS and the test user, and the \( i \)th
interfering BSs and its associated users. The methodology to
classify the characteristic of the equivalent channel gains
are given in the following steps:

1) **SNR characterization:** \( \tilde{g}_o \) is first characterized by
projecting the signal of the intended data-stream on
the null-space of the signals of the other data streams
that are multiplexed by the intended BSs. Note that we
may manipulate the resultant SNR such that the noise
variance is not affected by any random variable as in (4)
and the projection effect is contained in \( \tilde{g}_o \) and \( \tilde{g}_i \) only.

2) **Per-stream equivalent channel gain representation:**
\( \tilde{g}_i \) from each interfering BS is characterized based on
the manipulation done in the SNR characterization in
the previous step and characterizing \( |\tilde{a}_{i\ell,k}^o|^2 \) given in (3).

Note that \( |\tilde{a}_{i\ell,k}^o|^2 \) is characterized based on \( \{ \tilde{\mathbf{H}} \} \) which
captures the channel gain matrices involved in preceding
the signal at the \( i \)th BS and combining the interfering
symbols at the test UE.

Based on the aforementioned two steps, the equivalent
SISO-SINR given in Proposition 1 for the MIMO schemes
given in Table I is illustrated in this section.

1) **Single-Input-Multiple-Output (SIMO) systems:** for a
SIMO system, receive diversity is achieved using one transmit
antenna (i.e., \( L = N_i = 1 \)) and \( N_r \) receive antennas.
Since \( N_i = 1 \), then the intended and interfering channel
vectors are denoted by \( \mathbf{h}_o \) and \( \mathbf{h}_i \in \mathbb{C}^{N_r \times 1} \), respectively. By
employing Maximum Ratio-Combining (MRC) strategy to
combine the received signals, then \( \mathbf{w}_i = \mathbf{h}_i^H \). The equivalent
SISO channel gains are given by the following lemma.

**Lemma 3:** For a receive diversity SIMO setup technique,
the Gamma distribution parameters for the equivalent
intended and interfering channel gains are given by \( \tilde{\mathbf{H}}_o = \mathbf{W}_o \)
and \( \til(\til{\mathbf{H}_i} = \mathbf{W}_i \mathbf{H}_i \mathbf{V}_i \).

\[ \mathbf{W}_o \mathbf{H}_o \mathbf{V}_o \] where \( \mathbf{H}_o \) is the effective intended channel matrix depending on the employed
orthogonal code [4], [16]. Since no precoding is applied then 
\( V_o = V_i = I_{N_t} \). The equivalent SISO channel gains are given by the following lemma.

**Lemma 4:** A space-time encoder is employed at the network BSs. Then, the Gamma distribution parameters are given as 
\( m_o = N_u N_r \) and \( m_i = N_o \).

**Proof:** See Appendix D \( \blacksquare \)

3) **Zero-Forcing beamforming with ML Receiver (ZF-Rx):** ZF is a low-complexity suboptimal, yet efficient, technique to suppress interference from other transmitted symbols in the network. In order to recover the distinct transmitted streams, the received signal is multiplied by the equalizing matrix \( W_o = (H_o^H H_o)^{-1} H_o^H \) representing the pseudo-inverse of the intended channel matrix \( H_o \), whereas we assume no precoding at the transmitters side, i.e., \( V_o = V_i = I_{N_t} \). The equivalent SISO channel gains are given by the following lemma.

**Lemma 5:** By employing a ZF-Rx such that \( L = N_t \) distinct streams are being transmitted from the BSs, it can be shown that \( m_o = N_u N_r - N_t + 1 \) and \( m_i = N_t \).

**Proof:** See Appendix D \( \blacksquare \)

4) **Space-Division Multiple Access (SDMA):** SDMA is used to accommodate more users on the same resources to increase the network capacity. In this case, we consider that each BS is equipped by \( N_t \) transmit antennas and applies ZF transmission to simultaneously serve \( K \) single-antenna UEs that are independently and randomly distributed within its coverage area. To avoid rank-deficiency, we let \( N_t \geq K \), and hence, the number of data streams \( L = K \). A ZF-precoding in the form of \( V_o = [v_1, v_2, \ldots, v_K] \) such that \( v_i = q_i \) and \( q_i \) is defined as the \( l^{th} \) column of \( Q = H_o^{-1} (H_o H_o^{-1}) \) is applied by the test BS and no combining is applied at the single antenna test UE, and hence, \( W_o = I_K \). The interfering BSs apply the same precoding and combining strategy, and hence, the interfering precoding matrices are in the form \( V_i = [v_{i,1}, v_{i,2}, \ldots, v_{i,K}] \) such that \( v_{i,k} = \frac{q_{i,k}}{||q_{i,k}||} \) and \( q_{i,k} \) is the \( l^{th} \) column of \( Q_i = H_i^{-1} (H_i H_i^{-1}) \), note that \( H_i \neq H_o \) is the interfering channel matrix towards the corresponding intended users. The equivalent SISO channel gains are given by the following lemma.

**Lemma 6:** In a multi-user MIMO setup, the corresponding Gamma distribution parameters are given by \( m_o = N_t - K + 1 \) and \( m_i \approx K \).

**Proof:** See Appendix D \( \blacksquare \)

**Corollary 1:** **Single-User Beamforming (SU-BF):**

The SDMA scenario reduces to SU-MISO (i.e., transmit diversity) setting if the number of served users in the network is \( K = 1 \). Hence, \( m_o = N_t \) and \( m_i = 1 \).

5) **Spatially Multiplexed MIMO (SM-MIMO) systems:** for the sake of completeness, we also consider a spatially multiplexed MIMO setup with optimum joint maximum likelihood receiver. This case is important because it represents the benchmark for ZF decoding. Note that the analysis in this case is slightly different from the aforementioned schemes since joint detection is employed. Nevertheless, it can be represented via the equivalent SISO-SINR given in Proposition 1. Due to joint detection, no precoding/combining is applied such that \( W_o = V_o = V_i = I_{N_t} \). To analyze this case, we define the interference vector \( e(s, \hat{s}) = s - \hat{s} \) as the distance between \( s \) and \( \hat{s} \) and hence we derive the APEP, which is then used to approximate the ASEP as shown in the following lemma.

**Lemma 7:** For a SM-MIMO transmission, the Gamma distribution parameter for the equivalent intended channel gains is given by \( m_o = N_r \), while for the equivalent interfering links is given by \( m_i = N_t \). Furthermore, the averaged PEP over the distance distribution of \( \tau_o \) is

\[
\frac{\text{APEP}(||e||)}{\text{APEP}(||e||)} \approx 1 - \frac{1}{2} \left( 1 + \frac{m_o - 1}{m_i - 1} \right) \int_0^\infty \int_0^\infty 2\pi \lambda_{B} e^{-\pi \lambda_{B} z^2} \times \frac{1}{\sqrt{2}} e^{-\frac{(1 + \frac{m_o - 1}{m_i - 1})}{2}} F_1 \left( 1 - \frac{m_o}{2} \right) \times L_{|z'|} \left( \frac{m_o - 1}{4||e||^2} P_{e}^{-1} \right) dxdz. \]

(16)

Consequently, using the nearest neighbor approximation [48], where there are \( M \) equiprobable symbols, then

\[
\text{ASEP} \approx N_{||e||} \text{APEP}(||e||), \quad (17)
\]

where \( N_{||e||} \) is the number of constellation points having the minimum Euclidean distance denoted by \( \min_{s, \hat{s}} ||e(s, \hat{s})|| \) among all possible pairs of transmitted symbols, and hence is a modulation-specific parameter.

**Proof:** See Appendix D \( \blacksquare \)

VI. **NUMERICAL AND SIMULATION RESULTS**

In this section, we verify the validity and accuracy of the proposed unified model and discuss the potential of such unified framework for designing cellular networks. The simulations setup is as follows. The BSs transmit powers \( P \) vary while \( N_t \) is kept constant to vary the transmit SNR, the path-loss exponent \( \eta = 4 \), the noise power \( N_o = -90 \text{ dBm} \), the BSs intensity \( \lambda_B = 10 \text{ BSs/km}^2 \), \( \lambda_u = 20 \text{ users/km}^2 \), and finally the activity factor \( p = 1 \). The transmitted symbols are modulated using \( M \)-QAM modulation scheme.
A. Proposed model validation

We validate Theorem 1 for the derived ASEP and outage probability expressions via Monte-Carlo simulations, in Fig. 1, for a fixed number of antennas in order to seek a fair comparison among the MIMO schemes. That is, we consider different MIMO configurations with \( N_r = 2 \) and \( N_t = 2 \). Note that, for SIMO and MISO, \( N_t \) and \( N_r \) are set to 1, respectively. Further, in SDMA scenario, the number of single-antenna users served in the network is \( K = 2 \). The figure verifies the accuracy of the Gaussian signaling approximation and the developed ASEP and outage probability model, in which the analytic expressions perfectly match the simulations.

Fig. 2 validates Theorem 2 for the outage before and after retransmission against Monte-Carlo simulations. Fig. 2(a) shows the time diversity loss due to interference temporal correlation when compared to the independent interference scenario. The figure shows that assuming independent interference across time is too optimistic, since those UEs requiring retransmissions are then biased to ones where there are interferers nearby due to interference temporal correlation. Nevertheless, it is possible for the network operators to exploit more diversity in the second transmission to compensate for the expected degraded retransmission performance. Fig. 2(b) shows the effect of incremental diversity in the second transmission on the outage performance for \( m_o = 2 \) and \( m_t = 2 \). The figure shows that adjusting the MIMO configuration such that \( m_o = 5 \) in the retransmission compensates for the temporal correlation effect and achieves the same performance as independent transmission (e.g., up to 3 dB SIR improvement can be achieved).

<table>
<thead>
<tr>
<th>MIMO Setup</th>
<th>( m_o )</th>
<th>( m_t )</th>
<th>Ergodic Rate (bits/sec/Hz)</th>
<th>No. of bits per 4-QAM 16-QAM</th>
</tr>
</thead>
<tbody>
<tr>
<td>SIMO</td>
<td>2</td>
<td>1</td>
<td>2.9523</td>
<td>1.6926 2.1712</td>
</tr>
<tr>
<td>OSTBC</td>
<td>4</td>
<td>2</td>
<td>2.9771</td>
<td>1.7225 2.2044</td>
</tr>
<tr>
<td>ZF-Rx</td>
<td>1</td>
<td>2</td>
<td>3.1644</td>
<td>2.6300 2.7008</td>
</tr>
<tr>
<td>SDMA</td>
<td>1</td>
<td>2</td>
<td>3.1644</td>
<td>2.6300 2.7008</td>
</tr>
<tr>
<td>MISO</td>
<td>2</td>
<td>1</td>
<td>2.9523</td>
<td>1.6926 2.1712</td>
</tr>
<tr>
<td>SISO</td>
<td>1</td>
<td>1</td>
<td>1.4889</td>
<td>1.4780 1.6936</td>
</tr>
</tbody>
</table>

TABLE II: Overall achievable and actual rates gains per cell, w.r.t SISO networks, for the different MIMO setups, in an interference-limited scenario for \( M = 4, 16\)-QAM modulation scheme.

B. Diversity-Multiplexing Tradeoffs & Design Guidelines

Using similar parameters as Fig. 1, the results in Table II compare the performance of the considered MIMO configurations in terms of error probability, outage probability, ergodic rate, and throughput,\(^9\) to quantify the achievable gains w.r.t the SISO configuration. The results in both Fig. 1 and Table II clearly show the diversity-multiplexing tradeoff in cellular networks. The results show the outage probability improvement due to diversity, in which the OSTBC achieves the highest outage probability reduction. This is because OSTBC provides both transmit and receive diversity while MISO and SIMO provide either transmit or receive diversity. Note that despite that MISO and SIMO have the same performance, the SIMO is preferred because it relies on the receive CSI which is easier to obtain than the transmit CSI. The results also show the negative impact of multiplexing on the per-stream ASEP and outage probability in ZF-Rx and MU-MIMO schemes. However, multiplexing several streams per BS improves the overall ergodic rate and per-cell throughput as shown in Table II.

The results in Fig. 1 and Table II show the diversity-multiplexing tradeoffs that can be achieved for a \( 2 \times 2 \) MIMO setting. However, as \( N_t \) and \( N_r \) grow, several diversity and multiplexing tradeoffs are no longer straightforward to compare. Hence, it is beneficial to have a unified methodology to select the appropriate diversity, multiplexing, and number of antennas to meet a certain design objective. From Proposition 1 and the subsequent results we noticed two important insights: (i) The performances of MIMO schemes differ according to their relative \( m_o \) and \( m_t \) values. In other words, MIMO configurations with equal \( \frac{m_o}{m_t} \) ratio have approximately equivalent per-stream performance as shown in Fig. 3(a). Moreover, such equivalence in performance can be further verified by the numerical results of [16] where similar ASEP performance among different MIMO schemes having the same fading parameters has been reported. (ii) Multiplexing more data streams increases \( m_o \) and does not affect \( m_t \). On the other hand, diversity increases \( m_o \) and does not affect \( m_t \). In other words, \( m_o \) represents the diversity gain and \( m_t \) represents the number of independently multiplexed data streams per BS (i.e., multiplexing gain).

Based on the aforementioned insights, we plot the unified MIMO outage probability and ASEP performance results in Fig. 3. The Figs. 3(b) and 3(c) show the ASEP and outage probability for a varying ratio of \( \frac{m_o}{m_t} \) which can be used for all considered MIMO schemes. Conversely, Fig. 3 presents a unified design methodology for MIMO cellular networks as shown in Fig. 4. Such unified design provides reliable guidelines for network designers and defines the different flavors of the considered MIMO configurations in terms of achievable diversity and/or multiplexing gains. For instance, for an ASEP or outage probability constraint, the corresponding ratio \( \frac{m_o}{m_t} \) and modulation scheme are determined. Then, the network designer can determine the MIMO technique depending on the number of data streams (or number of users) that need to be simultaneously served (i.e., determine \( L \) or \( K \)). Finally, the number of transmit and receive antennas for the selected MIMO scheme can be determined from Table I. Figs. 3(b) and 3(c) clearly show that incrementing the ratio \( \frac{m_o}{m_t} \) enhances the diversity gain whereas decrementing it provides a higher multiplexing gain.

That is, network designers are able to maintain the same per-stream ASEP/outage probability by appropriately adjusting the operational parameters, namely, \( N_r, N_t \) and \( L \) (or \( K \) for SDMA). This is done by compensating \( m_t \) with the adequate \( m_o \) such that \( \frac{m_o}{m_t} \) is kept constant. For instance, consider a network that needs to increase the number of served users \( K \) without compromising the reliability performance of each served user. According to Table I and Fig. 3, this is achieved by keeping \( \frac{m_o}{m_t} = c \), where \( c \) is a constant, which hence costs the network additional \( \lceil K(c+1) - 1 \rceil \) transmit antennas per BS.

It is worth mentioning that a design based on the ASEP is more tangible as it is sensitive to the used modulation scheme and constellation size, as opposed to the outage
Fig. 1: ASEP and Outage probability performance validation for the different MIMO setups using the same number of antennas $N_t = 2$ and $N_r = 2$, at $p = 1$. Lines represent the proposed analysis and markers represent Monte-Carlo simulations.

(a) ASEP for 4-QAM and 16-QAM.

(b) Outage Probability versus SIR threshold $\theta$.

Fig. 2: ASEP and Outage probability performance validation. Lines represent the proposed analysis and markers represent Monte-Carlo simulations.

(a) The effect of interference correlation for different $m_o$ and $m_i$.

(b) Incremental diversity for the same inter-cell interference.

Fig. 3: Unified performance versus the ratio $\frac{m_o}{m_i}$ for an arbitrary MIMO setup.

(a) Outage Probability for $\frac{m_o}{m_i} = \frac{1}{2}$.

(b) Outage Probability for $\theta = 0, 5, 10$ dB.

(c) ASEP for 4-QAM and 16-QAM.
probability as shown in Fig. 1, and Table II. Also, note that increasing \( m_o \) for a fixed \( \frac{m_o}{m_i} \) ratio can slightly vary the outage probability due to the channel hardening effect as shown in Fig 3(a). However, such variation is shown to be negligible for \( m_o > 2 \). In fact, by direct inspection of eq. (18), it is clear that as \( m_o \) increases, the value of the summation also increases and therefore, the overall outage performance decreases, since the increase in \( m_o \) is interpreted as an enhancement in the desired signal. Nevertheless, if the improvement in the desired signal, i.e., \( m_o \), is compensated by an equivalent increase in the interfering signals, i.e., \( m_i \), the performance eventually saturates and no better performance can be achieved as long as the ratio \( \frac{m_o}{m_i} \) is kept constant. Another noteworthy observation is that the second term in (10), which corresponds to the \( \text{erfc}(\cdot) \) term in (5), requires threefold nested integrals that involve hypergeometric functions to evaluate the ASEP. Such integration is computationally complex to evaluate and may impose some numerical instability specially for large arguments of \( m_o \) and \( m_i \). In order to overcome such complexity and numerical instability, we invoke Jensen’s inequality to the \( \text{erfc}(\cdot) \) term in (5). Hence, the ASEP function becomes ASEP (\( \mathcal{T} \)) \( \geq \) \( w_1 \mathbb{E} \left[ \text{erfc} \left( \sqrt{\mathcal{T}} \right) \right] + w_2 \mathbb{E} \left[ \text{erfc} \left( \sqrt{\mathcal{T}} \right) \right]^2 \), which reduces one integral from the second term of (10). Using Jensen’s inequality yields a stable and accurate approximation compared to (10) as shown in Figure 3(c), where the red curves represent the numerically unstable ASEP performance as arguments grow, while the black curves represent the Jensen’s inequality tight upper bounds.

### VII. CONCLUSION

This paper provides a unified tractable framework for studying symbol error probability, outage probability, ergodic rate, and throughput for downlink cellular networks with different MIMO configurations. The developed model also captures the effect of temporal interference correlation on the outage probability after signal retransmission. The unified analysis is achieved by Gaussian signaling approximation along with an equivalent SISO-SINR representation for the considered MIMO schemes. The accuracy of the proposed model is verified against Monte-Carlo simulations. To this end, we shed lights on the diversity loss due to temporal interference correlation and discuss the diversity-multiplexing tradeoff imposed by MIMO configurations. Finally, we propose a unified design methodology to choose the appropriate diversity, multiplexing, and number of antennas to meet a certain design objective.

### APPENDIX

#### A: PROOF OF THEOREM 1

The ASEP expression in (10) is obtained by taking the expectation over \( \mathcal{T} \) and then using expressions from [49, eq. (11), (21)] as has been detailed in [29].

For the outage probability, conditioned on \( r_o \),

\[
\mathcal{O} (r_o, \theta) = \mathbb{E} \left[ P \left( g_o < \frac{\theta T}{P_r^o \eta} \right) \right]
\]

\[
= \mathbb{E} \left[ 1 - \sum_{j=0}^{m_o-1} \frac{1}{2} \left( \frac{\theta T}{P_r^o \eta} \right)^j \exp \left( \frac{-\theta T}{P_r^o \eta} \right) \right], \tag{18}
\]

where \( (c) \) follows from the CDF of the gamma distribution, and then (11) is obtained from the rules of differentiation of the LT, together with averaging over the PDF of \( r_o \).

Ergodic rate expression in (12) follows from [50, Lemma 1], and by exploiting the independence between the useful and interfering signals, as well as incorporating the CDF of the gamma random variable.

#### B: PROOF OF LEMMA 2

Let \( \tilde{\Psi}_1 \subset \Psi^o \) and \( \tilde{\Psi}_2 \subset \tilde{\Psi}^o \) be the sets of interfering BSs in the first and second time slots of transmissions, respectively. Exploiting the independent transmission assumption per time slot, \( \tilde{\Psi}_1 \) and \( \tilde{\Psi}_2 \) can be decomposed into three independent PPPs \( \{ \tilde{\Psi}^o \cap \tilde{\Psi}_1 \}, \{ \tilde{\Psi}^o \cap \tilde{\Psi}_2 \}, \{ \tilde{\Psi}^o \cap \tilde{\Psi}_1 \cap \tilde{\Psi}_2 \} \) with intensities \( p(1-p) \lambda_B, (1-p)\lambda_B \), and \( p^2 \lambda_B \), respectively. Substituting \( p \lambda_B = \lambda \), the joint LT of the two random variables \( T_1 \) and \( T_2 \) is given as shown in (19), where \( (d) \) is obtained from the PGFL and exploiting the independence between the PPPs \( \{ \tilde{\Psi}^o \cap \tilde{\Psi}_1 \}, \{ \tilde{\Psi}^o \cap \tilde{\Psi}_2 \}, \{ \tilde{\Psi}^o \cap \tilde{\Psi}_1 \cap \tilde{\Psi}_2 \} \) [35], and \( (e) \) follows from the LT of the two independent gamma distributed random variables \( \tilde{g_i}^{(1)} \) and \( \tilde{g_i}^{(2)} \). Solving the integral completes the proof.

#### C: PROOF OF THEOREM 2

The joint complementary cumulative distribution function (CCDF) of \( T_1 \) and \( T_2 \) is given by

\[
\Pr \left( T_1 > \theta, T_2 > \theta \right) = \mathbb{E} \left[ P \left( \tilde{g_1}^{(1)} > \frac{\theta T_1}{P_r^o \eta}, \tilde{g_2}^{(2)} > \frac{\theta T_2}{P_r^o \eta} \right) \right]
\]

\[
= \mathbb{E} \left[ \sum_{j=1}^{m_o-1} \sum_{j_2=0}^{m_o-1-j} \frac{1}{j! j_2!} \left( \frac{\theta}{P_r^o \eta} \right)^{j+j_2} \left( T_1^{j_1} T_2^{j_2} \right) \times \exp \left\{ -\frac{\theta (T_1 + T_2)}{P_r^o \eta} \right\} \right]
\]

\[
= \mathbb{E} \left[ \sum_{j=1}^{m_o-1} \sum_{j_2=0}^{m_o-1-j} \frac{1}{j! j_2!} \left( \frac{\theta}{P_r^o \eta} \right)^{j+j_2} \right.
\]

\[
\left. \times \left( \frac{\tilde{g_1}^{(j_1+j_2)}}{\tilde{g_1}^{(1)} \tilde{g_2}^{(2)}} \right) \mathcal{L}_{T_1, T_2}(\frac{\theta}{P_r^o \eta}, j_1, j_2) \right|_{j_1+j_2 = \frac{\theta}{P_r^o \eta}} \right]\ . \tag{20}
\]
such that (i) follows from the independence of $\tilde{y}_0^{(1)}$ and $\tilde{y}_0^{(2)}$ along with the CCDF of their Gamma distributions. (ii) is obtained by utilizing the LT identity $\int_0^1 t f(t, t_2) dt \leftrightarrow \partial^j/\partial z^j L(t_1, t_2)$, which can be proved as follows. First, we write the joint LT of two variables $t_1$ and $t_2$ as

$$L_{t_1, t_2}(z_1, z_2) = \int_0^\infty \int_0^\infty f(t_1, t_2) e^{-z_1 t_1} e^{-z_2 t_2} dt_1 dt_2,$$

then,

$$\frac{\partial^{j_1+j_2} L_{t_1, t_2}(z_1, z_2)}{\partial z_1^{j_1} \partial z_2^{j_2}} = \int_0^\infty \int_0^\infty f(t_1, t_2) e^{-z_1 t_1} e^{-z_2 t_2} dt_1 dt_2 = \int_0^\infty \int_0^\infty \frac{(-1)^{j_1+j_2} (t_1^{j_1} t_2^{j_2}) f(t_1, t_2) e^{-z_1 t_1} e^{-z_2 t_2} dt_1 dt_2}{t_1^{j_1} t_2^{j_2}}.$$ 

where the second equality follows by Leibniz rule and applying the rules of partial differentiation, which proves the identity.

**D: EQUIVALENT SISO MODEL PROOFS**

**Proof of Lemma 3**: In SIMO transmission, by applying MRC at the receiver side, for $\tilde{w}_0^H = h_0^H$, then the post-processed signal is given as

$$\tilde{y} = \tilde{w}_0^T \tilde{y} = \sqrt{P_{\tilde{r}_0}} \frac{1}{\epsilon} \| h_0 \|^2 s_0 + \sum_{r_i \in \Phi^o} \sqrt{P_{\tilde{r}_i}} h_i^H h_0^H s_i + h_0^H n. \quad (23)$$

We start with computing the effective noise variance since a post-processor is applied. The noise power is expressed as

$$\text{Var}_n \| h_0^H n \|^2 = N_0 \| h_0 \|^2. \quad (24)$$

Therefore, the random variable $\epsilon = \| h_0 \|^2$, is used to normalize the resultant interference power. The effective interference variance conditioned on the network geometry and the intended channel gains w.r.t $\tilde{s}_i$ is given by

$$I = \frac{1}{\epsilon} \text{Var}_s_i \sum_{r_i \in \Phi^o} \sqrt{P_{\tilde{r}_i}} \| h_i^H h_0^H s_i \|^2 = \sum_{r_i \in \Phi^o} P_{\tilde{r}_i} \| h_i^H h_0^H \|^2. \quad (25)$$

By inspection of the interference variance, it is clear that $\{H\} = \{H_0\}$. Also, we notice that there exists only one coefficient $\alpha^{(i)}_{i,k} = h_i^H h_i^H$. Recall that the number of independent coefficients $\alpha^{(i)}_{i,k}$ depends on the number of independent transmitted streams, which is equal to one in the SIMO case. Accordingly, $\tilde{y}_i = |\alpha^{(i)}_{i,k}|^2 \sim \text{Gamma}(m_i, 1)$, with $m_i = 1$. Similarly, conditioned on the intended and interfering channel gains, the received signal power, w.r.t the transmitted signal, can be shown to be

$$S = \frac{1}{\epsilon} \text{Var}_s \left[ \sqrt{P_{\tilde{r}_0}} \| h_0 \|^2 s_0 \right] = \frac{P_{\tilde{r}_0} \| h_0 \|^2}{\epsilon}. \quad (26)$$

Therefore, $\tilde{y}_0 = \| h_0 \|^2 \sim \text{Gamma}(m_0, 1)$ where $m_0 = N_r$.

**Proof of Lemma 4**: Employing OSTBC, the received vector at a typical user at time instant $\tau$, $N_t \leq T$, is given by

$$y(\tau) = \sqrt{P_{\tilde{r}_0}} s + \sum_{r_i \in \Phi^o} \sqrt{P_{\tilde{r}_i}} \tilde{h} s_i + n(\tau). \quad (27)$$

Let $\mathcal{Y}$ be the stacked vector of received symbols over $T$ intervals, and let $L = N_t$, such that

$$\mathcal{Y} = \sqrt{P_{\tilde{r}_0}} \frac{1}{\epsilon} \mathbf{H}_{\text{eff}} s + \mathbf{i}_{\text{agg}} + n. \quad (28)$$

where $\mathcal{Y} \in \mathbb{C}^{T \times N_r}$, and $i_{\text{agg}}$ is the concatenated aggregate interference $T \cdot N_r \times 1$ vector. The effective channel matrix $\mathbf{H}_{\text{eff}} \in \mathbb{C}^{T \times N_r \times N_t}$ is expressed as a linear combination of the set of dispersion matrices $A$ and $B$ chosen according to the adopted orthogonal space-time code as follows [3], [51].

$$\mathbf{H}_{\text{eff}} = \sum_{j=1}^{N_x} \sum_{j=1}^{N_y} \alpha_{j_q} A_{j_q} + j \beta_{j_q} B_{j_q}. \quad (29)$$

where $h_{j_q} = \alpha_{j_q} + j \beta_{j_q}$. Moreover, $\| \mathbf{H}_{\text{eff}} \|_2^2 = \| \mathbf{H}_0 \|_2^2$ is the squared Frobenius norm of the intended channel matrix. Hence, $\| \mathbf{H}_{\text{eff}} \|_2^2 \sim \frac{\lambda}{2} \| \mathbf{H}_0 \|_2^2$ [1]. Moreover, the aggregate interfering signals are expressed as

$$\mathbf{i}_{\text{agg}} = \sum_{r_i \in \Phi^o} \sqrt{P_{\tilde{r}_i}} \tilde{h}_i \tilde{s}_i. \quad (30)$$

such that $\mathbf{H}_{\text{eff}}$ is defined similar to (29). For detection, we equalize the effective channel matrix at the receiver side by $\mathbf{W}_0$. Hence, the received vector $\tilde{\mathcal{Y}}$ is written as

$$\tilde{\mathcal{Y}} = \mathbf{W}_0 \mathcal{Y} = \sqrt{P_{\tilde{r}_0}} \frac{1}{\epsilon} \| h_0 \|^2 s + \sum_{r_i \in \Phi^o} \sqrt{P_{\tilde{r}_i}} \tilde{A}_i \tilde{s}_i + \mathbf{w}. \quad (31)$$

such that $\mathbf{w} = \mathbf{W}_0 n$ and $A_i = \mathbf{W}_0 \mathbf{H}_{\text{eff}}$ with elements $a_{i,k}^{(i)}$ as defined in (2). Without loss of generality, we consider the detection of the $l^{th}$ arbitrary symbol from the received vector $\tilde{\mathcal{Y}}$. Due to the adopted Gaussian signaling scheme, we lump interference with noise, and thus it is essential to obtain the interference variance. First, let us define $q_{i,k}$ as the $k^{th}$ column of the matrix $\mathbf{H}_{\text{eff}}$, similarly, $q_{i,k}$ is the $k^{th}$ column of the matrix $\mathbf{H}_{\text{eff}}$. Then, the received interference variance for...
the $l^{th}$ symbol denoted as $I_l$, computed w.r.t the interfering symbols $\tilde{s}_i$, can be derived as

$$I_l = \text{Var}_{\tilde{s}_i} \left[ \sum_{r_i \in \Phi^o} \sum_{t_i=1}^{N_t} P_{r_i}^{-\eta} \frac{\lVert g_t q_{t,k} \lVert H_i}{H_i} \tilde{s}_i, k \right]$$

$$= \sum_{r_i \in \Phi^o} P_{r_i}^{-\eta} \frac{\lVert g_t q_{t,k} \lVert H_i}{H_i} \text{Var}_{\tilde{s}_i} \left[ \tilde{s}_i, k \right]$$

(32)

where the summation is over the $N_s$ active antennas per transmission. Note that, conditioned on $\tilde{H} = \{H_o\}$, $q_{i,k} = \frac{g^{H}_t a_{i,k}}{H_i}$ is a normalized and independently weighted sum of complex Gaussian random variables, thus $a_{i,k} \sim \mathcal{CN}(0, 1)$. Although a post-processor is applied, the noise variance is maintained to be $N_o$. Thus, $\tilde{g}_t \sim (m_i, \Omega_i)$ with $m_i = N_o$ and $\Omega_i = 1$. Similarly, the received signal power is found to be

$$S = \text{Var}_{\tilde{s}_i} \left[ \sqrt{P_{r_i}^{-\eta}} \lVert H_i \lVert \tilde{s}_i \right] = P_{r_i}^{-\eta} \lVert H_i \lVert^2$$

(33)

Proof of Lemma 5: Without loss of generality, we focus on the detection of an arbitrary symbol $l$ from the received vector $\tilde{y} = \mathbf{W}_o \tilde{s}_i$, given by

$$\tilde{y}_l = \sqrt{P_{r_i}^{-\eta}} s_l + \sum_{r_i \in \Phi^o} \sqrt{P_{r_i}^{-\eta}} \mathbf{W}_o H_i s_i + \mathbf{W}_o n_i$$

(34)

which is similar to (2). First we need to to obtain the received noise variance since a post-processing matrix is applied and thus the noise variance is scaled. Conditioned on $H_o$, the received noise power is defined as

$$\text{Var}_{n} \left[ \mathbf{w}^T_{o,l} \mathbf{n} \right] = \mathbf{w}^T_{o,l} \mathbb{E} \left[ \mathbf{n} \mathbf{n}^H \right] \mathbf{w}^*_{o,l} = N_o \left( \mathbf{W}_o \mathbf{W}_o^H \right)_{ll}^{-1}$$

(35)

Then, the scaling random variable is $\epsilon = \left( \mathbf{H}_o \mathbf{H}_o^H \right)_{ll}^{-1}$. Next, we obtain the effective interference variance from the $l^{th}$ received symbol as

$$I_l = \frac{1}{\epsilon} \text{Var}_{\tilde{s}_i} \left[ \tilde{y}_l, l \right] = \frac{1}{\epsilon} \sum_{r_i \in \Phi^o} P_{r_i}^{-\eta} \left( \mathbf{W}_o \mathbf{W}_o^H \right)_{ll} \left( \mathbf{H}_o \mathbf{H}_o^H \right)_{ll}^{-1} \sum_{r_i \in \Phi^o} P_{r_i}^{-\eta} \left( \mathbf{H}_o \mathbf{H}_o^H \right)_{ll}^{-1}$$

(36)

The processing resulting interference channel set $\{ \tilde{H} \} = \emptyset$. Therefore, $a_{i,k} = \left( \mathbf{H}_o \mathbf{H}_o^H \right)_{ll}$ and $\tilde{g}_t \sim (m_i, \Omega_i)$, with $m_i = N_t$ and $\Omega_i = 1$. The received signal power is similarly computed as

$$S = \frac{1}{\epsilon} \text{Var}_{\tilde{s}_i} \left[ \tilde{y}_l, l \right] = \frac{P_{r_i}^{-\eta}}{\left( \mathbf{H}_o \mathbf{H}_o^H \right)_{ll}^{-1}}$$

(37)

Since $\tilde{g}_o = \left( \mathbf{H}_o \mathbf{H}_o^H \right)_{ll}^{-1} \sim \text{Inv-Gamma}(N_o - N_t + 1, 1)$ [1]. Then, we can let $\frac{1}{\epsilon} = \tilde{g}_o \sim \text{Gamma}(m_o, \Omega_o)$, where $m_o = N_r - N_t + 1$ and $\Omega_o = 1$.

Proof of Lemma 6: In a multi-user MIMO setting, we introduce a slight abuse of notation for the intended and interfering channel matrices such that they are of dimensions $K \times N_t$. The received interference power at user $l$ where $1 \leq l \leq K$, averaged over the interfering symbols $\tilde{s}_i$ is given by

$$I_l = \text{Var}_{\tilde{s}_i} \left[ \sum_{r_i \in \Phi^o} \sqrt{P_{r_i}^{-\eta}} h_{i,k} \tilde{s}_i \right] = \sum_{r_i \in \Phi^o} P_{r_i}^{-\eta} \left| \tilde{s}_i \right|^2$$

(38)

where $h_{i,l}$ is the $l^{th}$ row of $H_i$ and $\{ \tilde{H} \} = \{ \mathbf{V}_i \}$. Also, $\left| \mathbf{h}_i \mathbf{V}_i \right| = \sum_{l=1}^{K} \left| a_{i,l} \right|^2$. However, the column vectors of $\mathbf{V}_i$ are not independent. Therefore, conditioned on $\mathbf{v}_{i,l}$, the linear combination $\sum_{l=1}^{K} \left| a_{i,l} \right|^2$ does not follow a Gamma distribution. Nevertheless, for tractability we approximate this summation by a Gamma distribution. Thus, $\tilde{g}_t \sim \mathcal{G}(m_t, \Omega_t)$, where $m_t = K$ and $\Omega_t = 1$ by assuming such independence. This renders the aggregate interference power distribution at user $l$ an approximation. Similarly, the useful signal power at user $l$ is straightforward to be obtained, after appropriate diagonalization, as $S = \frac{1}{\epsilon} \tilde{g}_o \sim \mathcal{G}(m_o, \Omega_o)$, with $m_o = N_t - K + 1$ and $\Omega_o = 1$ [2]. This can also be interpreted as having the preceding matrix nulling out $K - 1$ directions out of the $N_t$ subspace at the transmitter side.

Proof of Lemma 7: Since, there are $N_t$ distinct multiplexed symbols to be transmitted, we will study the pairwise error probability (PEP) of two distinct transmitted codewords, denoted as $\mathcal{P}(\mathbf{e}) = \mathbb{P} \left[ \sum_{l=1}^{K} \left| a_{i,l} \right|^2 > \delta \right]$, where $\mathbf{I}_{agg} = \sum_{r_i \in \Phi^o} I_i$. Conditioned on the channel matrices $\mathbf{H}_o$ and $\mathbf{H}_i$, and considering the Gaussian signaling approximation, the L.H.S of the above inequality represents the interference-plus-noise power and is a Gaussian random variable, denoted as $\mathcal{V}$ with zero-mean and variance $\sigma^2_v$, thus,

$$\mathcal{P}(\mathbf{e}) = \frac{1}{2} \text{erfc} \left( \frac{\delta e^{H}_o \mathbf{H}_o e}{\sqrt{2\sigma^2_v}} \right)$$

where the variance $\sigma^2_v$ is given by

$$\sigma^2_v = 2 \left[ e^{H}_o \mathbf{H}_o^H \mathbf{H}_o e \right] \left( N_o + \sum_{r_i \in \Phi^o} P_{r_i}^{-\eta} \sum_{k=1}^{N_t} \frac{1}{\left| \left( \mathbf{H}_o \mathbf{H}_o^H \right)_{k,k} \right|^2} \right)$$

(41)

By following the same convention used in this paper, it is clear that the interference power is represented as

$$I_l = \sum_{r_i \in \Phi^o} P_{r_i}^{-\eta} \sum_{k=1}^{N_t} \frac{\left| \left( \mathbf{H}_o \mathbf{H}_o^H \right)_{k,k} \right|^2}{\left| \mathbf{H}_o \mathbf{H}_o^H e \right|^2}$$

(42)
\[ S = \mathbf{H} \mathbf{d}^H \mathbf{H}_o \mathbf{e} = \| \mathbf{e} \|^2 (\mathbf{H}^H \mathbf{H}_o)_{ll}, \] hence it is straightforward to see that \((\mathbf{H}^H \mathbf{H})_{ll}\) has a \(\chi^2(N_r)\) distribution. Thus, 
\[ \tilde{g}_o = (\mathbf{H}^H \mathbf{H}_o)_{ll} \sim \text{Gamma}(m_o, \Omega_o) \], with \(m_o = N_r\) and \(\Omega_o = 1\). Then, the conditional pairwise error probability is expressed by

\[ P(e) = \frac{1}{2} \text{erfc} \left( \sqrt{4 \left( N_r^2 + \sum_{r \in \Psi} P_{r_i}^{-\gamma} \tilde{g}_o \right)} \right) \]  

(43)

REFERENCES

main research interests include signal processing, stochastic geometry and research focuses on the use of stochastic geometry in cellular networks. Her


Laila Hesham Afify received her B.Sc. degree in Electrical Engineering from Cairo University, Egypt, in 2009. She worked as a research assistant at the Wireless Intelligent Network Center (WINC), Nile University, Giza, Egypt from 2009 to 2011. She received her M.Sc. degree in Wireless Communications from Nile University in 2011. Currently, she is pursuing her Ph.D. degree at King Abdullah University of Science and Technology (KAUST), Saudi Arabia. From 2011 to 2012, she was jointly affiliated with Nile University and the American University in Cairo as a junior scientist. She is also the recipient of the Best Paper Award in ICC 2015 workshop on Small Cells and 5G networks. Her research focuses on the use of stochastic geometry in cellular networks. Her main research interests include signal processing, stochastic geometry and cognitive radio systems.

Hesham ElSawy (S’10, M’14) received the B.Sc. degree in Electrical Engineering from Assiut University, Assiut, Egypt, in 2006, the M.Sc. degree in Electrical Engineering from Arab Academy for Science and Technology, Cairo, Egypt, in 2009, and the Ph.D. degree in Electrical Engineering from the University of Manitoba, Winnipeg, MB Canada, in 2014. Currently, he is a postdoctoral fellow with the Computer, Electrical, and Mathematical Sciences and Engineering Division, King Abdullah University of Science and Technology (KAUST), Saudi Arabia, and an adjunct member at the school of Computer Science & Engineering, York University, Canada. During the period of 2006-2010, he worked at the National Telecommunication Institute, Egypt, where he conducted professional training both at the national and international levels, as well as research on network planning. From 2010 to 2014, he worked with TRTech, Winnipeg, MB, Canada, as a Student Researcher. For his academic excellence, he has received several academic awards, including the NSERC Industrial Postgraduate Scholarship during the period of 2010-2013, and the TRTech Graduate Students Fellowship in the period of 2010-2014. He also received the best paper award in the ICC 2015 workshop on small cells and 5G networks. He is recognized as an exemplary reviewer by the IEEE Transactions of communication in 2015 & 2016. His research interests include statistical modeling of wireless networks, stochastic geometry, and queueing analysis for wireless communication networks.

Tareq Al-Naffouri received the B.S. degrees in mathematics and electrical engineering (with first honors) from King Fahd University of Petroleum and Minerals, Dhahran, Saudi Arabia, the M.S. degree in electrical engineering from the Georgia Institute of Technology, Atlanta, in 1998, and the Ph.D. degree in electrical engineering from Stanford University, Stanford, CA, in 2004. He was a visiting scholar at California Institute of Technology, Pasadena, CA, from January to August 2005 and during summer 2006. He was a Fulbright scholar at the University of Southern California from February to September 2008. He has held internship positions at NEC Research Labs, Tokyo, Japan, in 1998, Adaptive Systems Lab, University of California at Los Angeles in 1999, National Semiconductor, Santa Clara, CA, in 2001 and 2002, and Beceem Communications Santa Clara, CA, in 2004. He is currently an Associate at the Electrical Engineering Department, King Abdullah University of Science and Technology (KAUST). His research interests lie in the areas of sparse, adaptive, and statistical signal processing and their applications and in network information theory. He has over 160 publications in journal and conference proceedings, 9 standard contributions, 10 issued patents, and 6 pending.

Dr. Al-Naffouri is the recipient of the IEEE Education Society Chapter Achievement Award in 2008 and Al-Marai Award for innovative research in communication in 2009. Dr. Al-Naffouri has also been serving as an Associate Editor of Transactions on Signal Processing since August 2013.

Mohamed-Slim Alouini (S’94, M’98, SM’03, F’09) was born in Tunis, Tunisia. He received the Ph.D. degree in Electrical Engineering from the California Institute of Technology (Caltech), Pasadena, CA, USA, in 1998. He served as a faculty member in the University of Minnesota, Minneapolis, MN, USA, then in the Texas A&M University at Qatar, Education City, Doha, Qatar before joining King Abdullah University of Science and Technology (KAUST), Thuwal, Makkah Province, Saudi Arabia as a Professor of Electrical Engineering in 2009. His current research interests include the modeling, design, and performance analysis of wireless communication systems.