

On the Transmission Probabilities in Quantum Key Distribution Systems over FSO Links

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Abstract

In this paper, we investigate the transmission probabilities in three cases (depending only on the legitimate receiver, depending only the eavesdropper, and depending on both legitimate receiver and eavesdropper) in quantum key distribution (QKD) systems over free-space optical links. To be more realistic, we consider a generalized pointing error scenario, where the azimuth and elevation pointing error angles caused by stochastic jitters and vibrations in the legitimate receiver platform are independently distributed according to a non-identical normal distribution. Taking these assumptions into account, we derive approximate expressions of transmission probabilities by using the Gaussian quadrature method. To simplify the expressions and get some physical insights, some asymptotic analysis on the transmission probabilities is presented based on asymptotic expression for the generalized Marcum Q-function when the telescope gain at the legitimate receiver approaches to infinity. Moreover, from the asymptotic expression for the generalized Marcum Q-function, the asymptotic outage probability over Beckmann fading channels (a general channel model including Rayleigh, Rice, and Hoyt fading channels) can be also easily derived when the average signal-to-noise ratio is sufficiently large, which shows the diversity order and array gain.

Index Terms

Beckmann distribution, free-space optics, generalized pointing errors, quantum key distribution, and transmission probability.

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I. INTRODUCTION

Quantum communication provides a promising solution to break the Shannon channel capacity limit [1] and achieve an unprecedented level of security [2] simultaneously, two competing tasks which cannot be realized in conventional technologies [3]. In this context, quantum key distribution (QKD) or quantum cryptography is a method for sharing the secret cryptographic keys between two legitimate parties to achieve the secure communications by taking advantage of the laws of quantum mechanics and quantum non-cloning theorem [4], [5]. However, further investigation and real application of QKD did not attract much attention until it was proved that the quantum computer was able to break public-key cryptosystems, which are commonly used in the modern cryptography [6], [7].

The connection implementation of QKD includes two main medium, i.e., fiber cable and free-space optics (FSO). Compared to the fiber cable, the implementation of QKD over FSO links is more convenient and easier due to the flexibility of the free space connection and satellite support for distributing quantum keys worldwide [8], [9]. Moreover, FSO is an alternative transport technology to interconnect high capacity networking segments in current and future communication systems, because of its cost-effectiveness, high-bandwidth availability, and interference-immunity [10]. The authors in [11]–[18] have presented some basic performance analysis works over classical FSO links or hybrid RF-FSO links, but those previous works do not consider the QKD mechanism. Actually, the research work about QKD over FSO links in the communications field is considerably limited.

In practical systems, the security of QKD strongly depends on the device implementation [19]–[22]. That is, a third party may have a side channel by making use of any deviation of a QKD device from the theoretical model. For example, two zero-error attacks on commercial QKD systems were reported where the defects in quantum signal encoding and detection were exploited [19], [20]. Besides, some imperfections in QKD designs can be also exploited by a plethora of quantum hacking attacks using current technologies [23].

In real commercial QKD implementations, a single-photon mechanism is typically used to convey the information, and the corresponding common detection scheme is called single photon avalanche photodiodes (SPADs) where the SPAD diode is operated in Geiger mode (reverse-biased above the breakdown voltage to create an avalanche) to count single-photons [24]. However, this detection scheme can lead to the information leakage, because the avalanche created by the incoming photon can emit a secondary photon which may be intercepted by a third party, namely the eavesdropper. This secondary photon emission (the photon emission in the sender is the first emission) is called backflash, which is quenched along with the avalanche, i.e., the backflash is quenched if the detection bias is lowered

below the breakdown voltage [22], [24]. Previous measurements show that the probability of detecting backflash is greater than 0.4%, and more than 0.04 photons emerging from the devices are contained in the backflash given the 10% nominal detection efficiency of SPADs [22], [25]. These measurements provide a reference for the backflash resulting in the information leakage, although the measurement results may change in different detector types and optical components.

An unevadable vulnerability in the FSO QKD systems is the random pointing error due to stochastic jitters and vibrations which can be caused by building sway, thermal expansion, and week earthquakes, in the urban FSO systems [26]–[28]. Similarly, for satellite communications, there are internal and external reasons for stochastic vibrations [29]. For example, the structure deformations caused by temperature gradients, and the gravitational force inhomogeneity over the satellite orbit, are two main external reasons for stochastic vibrations in satellite systems. The internal sources include electronic noise, antenna pointing operation, and solar array driver [30].

The authors in [31] first investigated the performance of received powers at both the legitimate receiver and eavesdropper in FSO QKD systems with taking random pointing errors into account, and derived the closed-form expressions for the corresponding average received powers. However, the authors in [31] assumed that the azimuth and elevation pointing errors are identically independently distributed, and more specifically, these two pointing errors are modeled by Gaussian distribution with zero-mean and the same variance which may be a little ideal. In the practical systems, the mean and variance of these two pointing errors are typically different. Moreover, the authors in [31] did not consider the transmission probability depending on the received power threshold, which is very important and useful for the system evaluation and design. This is because we need to know the transmission probability depending on some conditions in the average level, apart from the average received powers, when evaluating and designing FSO QKD systems.

Actually, the pointing error angle, divided into the azimuth and elevation pointing errors, can be modeled by the Beckmann distribution [32]–[35], a generalized model including the Rayleigh, Hoyt and Rice models. Specifically, the Beckmann model is reduced into the Hoyt model for zero-mean and different variance of two sub-part pointing errors, and the different non-zero mean and equal variance case refers to the Rice case in the Beckmann distribution. The zero-mean and equal variance case considered in [31] is the most simple scenario in the Beckmann distribution, denoted by the Rayleigh case. The authors in [36] expanded the pointing error model in [31] to the Beckmann distribution. In [36], exact closed-form expressions for the average received powers at both the legitimate receiver and eavesdropper were derived, as well as finding the maximum points of the telescope gain at the legitimate receiver in some special cases analytically. However, the authors in [36] still did not investigate the transmission

probabilities depending on a variety practical conditions, which is also important for the FSO QKD system evaluation. Similar to [37]–[39], the transmission probability in this paper is defined as the probability that the received power is satisfied one or more pre-set thresholds in the FSO QKD system, which is obviously a natural variant of the outage probability in traditional communications [35].

Motivated by observing those facts outlined above, we investigate the performance of QKD systems over FSO links in terms of transmission probabilities depending on three different conditions. The main contributions of this paper are summarized as follows:

- 1) Closed-form expressions with a high accuracy for transmission probabilities depending on three different conditions, i.e., legitimate receiver, eavesdropper and both legitimate receiver and eavesdropper, are derived based on the Gaussian quadrature rule, where the accuracy grows with increasing the summation terms in the Gaussian quadrature.
- 2) Asymptotic expression for the generalized Marcum Q-function is derived after some mathematical manipulations, which can be used to derive the asymptotic outage probability over Beckmann fading channels when the average signal-to-noise ratio (SNR) approaches to infinity, showing the diversity order and array gain, since those two metrics govern the outage probability behaviour in high SNRs.
- 3) By using the asymptotic result for the generalized Marcum Q-function, the asymptotic expressions for three transmission probabilities are easily derived, which are valid in the high value region of the telescope gain at the legitimate receiver. Besides providing some insights, these asymptotic expressions are significantly concise, resulting in a much faster calculation than the analytical expressions that need to be computed based on the Gaussian quadrature rule.
- 4) We also present some specific expressions for those three transmission probabilities in some simplified cases which result in exact expressions or more concise forms. More specifically, exact closed-form expressions in Rayleigh, Hoyt and Rice cases (three special cases of the Beckmann distribution) for the transmission probability depending only on the legitimate receiver are given. In the Rayleigh case, we present a more concise expression for the transmission probability depending on both the legitimate receiver and eavesdropper.

The remainder of this paper is organized as follows. The system model is presented in Section II. The transmission probabilities depending on three different conditions are analyzed in Sections III, IV and V, respectively. In Section VI, some numerical results are generated and used to validate the correctness of derived closed-form expressions, as well as presenting some interesting comparisons. Section VII finally concludes the paper.

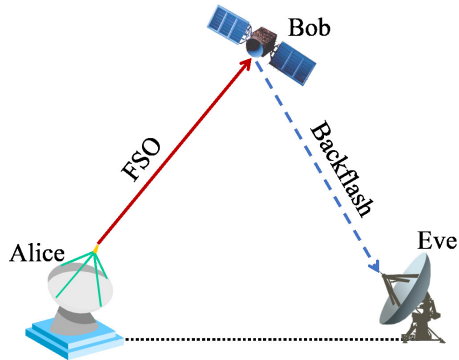


Fig. 1: Secure QKD System Over FSO links

II. SYSTEM MODEL

As depicted in Fig. 1, there is a sender (Alice) located on a absolutely static platform¹ communicating a legitimate receiver (Bob) located on a platform suffering from stochastic jitters and vibrations, such as a laser satellite system, over a FSO link in open areas. This vibrating platform in the legitimate receiver results in a random pointing error, where the stochastic deviation angle (θ) is divided into two parts, i.e., the azimuth pointing error (θ_H) and the elevation pointing error (θ_V), and therefore, θ can be written as

$$\theta = \sqrt{\theta_H^2 + \theta_V^2},$$

where θ_H and θ_V are normally assumed to be independent Gaussian random variables, i.e., $\theta_H \sim \mathcal{N}(\mu_H, \sigma_H^2)$ and $\theta_V \sim \mathcal{N}(\mu_V, \sigma_V^2)$, where μ_H and σ_H^2 (or μ_V and σ_V^2) represent the mean and variance of θ_H (or θ_V), respectively.

The SPADs detection scheme adopted by the legitimate receiver is assumed in this system setting. The received power at Bob is² [26], [29], [31]

$$P_D(\theta) = K_1 L(\theta) G_D,$$

where $L(\theta) = \exp(-G_D \theta^2)$ is the pointing loss factor, $G_D = (\pi d_D / \lambda_1)^2$ is the telescope gain at the legitimate receiver, d_D is the unobscured circular aperture diameter of the telescope, λ_1 is the wavelength,

¹We can also consider a non-static platform for the sender. In a relative motion aspect, if the sender is assumed to be relatively static to the receiver, this will induce the same analysis.

²This power can be also regarded as an average received power over instantaneous received photon counts influenced by both the shot noise and the dead time of the SPAD receiver. Here, we focus on the transmission probability depending on the average power, rather than the instantaneous performance.

and K_1 is a constant depending only on the system design, given by

$$K_1 = \eta_q P_S G_S \eta_S \eta_D \frac{L_A(Z_1)}{Z_1^2} \left(\frac{\lambda_1}{4\pi} \right)^2,$$

in which P_S is the optical transmitter power, η_q is the quantum efficiency, G_S is the telescope gain of the sender, $L_A(\cdot)$ is the atmospheric loss, and Z_1 is the distance between the sender and legitimate receiver.

As discussed in the introduction section, the SPADs detection scheme leads to the backflash due to the secondary photon emission caused by the avalanche. A third party (eavesdropper, Eve) can make use of this backflash to intercept the secondary photon, and thereby wiretapping the conveyed information from the sender to the legitimate receiver. The received power at the eavesdropper is given by [31], [36]

$$P_E(\theta_E, \alpha) = K_2 P_D(\theta) L(\theta_E) G_D,$$

where $\theta_E = \sqrt{(\theta_V + \alpha)^2 + \theta_H^2}$, α is the pointing direction error angle in the wiretap FSO link, and K_2 is a system constant, given by

$$K_2 = \eta_B \eta_q \eta_D \eta_E G_E \frac{L_A(Z_2)}{Z_2^2} \left(\frac{\lambda_2}{4\pi} \right)^2,$$

in which η_B is the probability of backflash, Z_2 is the distance between the legitimate receiver and eavesdropper, η_E and G_E are the optical efficiency and telescope gain of the eavesdropper respectively, and λ_2 is the backflash wavelength.

III. TRANSMISSION PROBABILITY DEPENDING ONLY ON LEGITIMATE RECEIVER

In this section, we want to evaluate the transmission probability performance given a received power threshold (λ_D) at the legitimate receiver. In this context, the transmission probability depending only on the legitimate receiver (TPLR) is defined as

$$\text{TPLR} = \Pr \left\{ P_D(\theta) \geq \lambda_D \right\} = \Pr \left\{ K_1 G_D L(\theta) \geq \lambda_D \right\}, \quad (1)$$

which can be further written by substituting the expression for $L(\theta)$, given by

$$\text{TPLR} = \Pr \left\{ \theta^2 \leq \frac{-1}{G_D} \ln \frac{\lambda_D}{K_1 G_D} \right\}. \quad (2)$$

Let $\Theta_D = \frac{-1}{G_D} \ln \frac{\lambda_D}{K_1 G_D}$. By substituting the probability density functions (PDFs) of θ_V and θ_H into (2), the TPLR for arbitrary μ_V , μ_H , σ_V and σ_H can be derived as

$$\begin{aligned} \text{TPLR} &= \Pr \left\{ \theta^2 \leq \Theta_D \right\} = \Pr \left\{ \sqrt{\theta_V^2 + \theta_H^2} \leq \sqrt{\Theta_D} \right\} \\ &= \iint_{\sqrt{\theta_V^2 + \theta_H^2} \leq \sqrt{\Theta_D}} f_{\theta_V}(\theta_V) f_{\theta_H}(\theta_H) d\theta_V d\theta_H = \iint_{\sqrt{\theta_V^2 + \theta_H^2} \leq \sqrt{\Theta_D}} \frac{\exp \left(-\frac{(\theta_V - \mu_V)^2}{2\sigma_V^2} - \frac{(\theta_H - \mu_H)^2}{2\sigma_H^2} \right)}{2\pi\sigma_V\sigma_H} d\theta_V d\theta_H, \quad (3) \end{aligned}$$

where the double integral, unfortunately, cannot be solved in a closed-form, and therefore, there are some approximation methods for this double integral, such as [32] and [33]. However, those approximation methods proposed by [32] and [33] are still complicated for calculation.

Here, we provide another approximation method based on the Gaussian quadrature rule [40, Ch. 9], shown in Theorem 1.

Theorem 1. *An approximate result for the TPLR based on the Gaussian quadrature rule is*

$$\text{TPLR} \approx 1 - \frac{1}{\sqrt{\pi}} \sum_{i=1}^N \omega_i f_{\text{TPLR}}(x_i), \quad (4)$$

where N , ω_i , and x_i are the summation terms, weights, and selected points of the Gauss-Hermite quadrature (GHQ, a special case of Gaussian quadrature), and $f_{\text{TPLR}}(\cdot)$ is given by

$$f_{\text{TPLR}}(x) = Q_{\frac{1}{2}} \left(\sqrt{\lambda}, \frac{1}{\sigma_H} \sqrt{\Theta_D - (\sqrt{2}\sigma_V x + \mu_V)^2} \times \mathbb{I} \left\{ \Theta_D \geq (\sqrt{2}\sigma_V x + \mu_V)^2 \right\} \right), \quad (5)$$

in which $\lambda = \mu_H^2 / \sigma_H^2$ represents the noncentrality parameter, $Q(\cdot, \cdot)$ denotes the generalized Marcum Q -function [42], and $\mathbb{I}\{\cdot\}$ denotes the indicator function, i.e., $\mathbb{I}\{\mathcal{A}\} = \begin{cases} 1, & \text{if } \mathcal{A} \text{ is true;} \\ 0, & \text{otherwise.} \end{cases}$

Proof. See Appendix A. □

Remark 1. *Although a high accuracy requires many terms in (4), resulting in a much slower calculation, especially when $Q_{\frac{1}{2}}(\cdot)$ cannot be directly calculated in some softwares, such as Matlab, Theorem 1 provides an analytical tool to investigate the TPLR. We will present an asymptotic expression presented in the III-A subsection, rather than (4), for getting a high accuracy result in a special case.*

As θ_V follows the Gaussian distribution, for $\sigma_V^2 \ll \sigma_H^2$, according to [44, Eq. (4)], the TPLR in (54) can be robustly approximated by

$$\begin{aligned} \text{TPLR} &= \mathbb{E}_{\theta_V} \left\{ 1 - Q_{\frac{1}{2}} \left(\sqrt{\lambda}, \sqrt{\frac{\Theta_D - \theta_V^2}{\sigma_H^2}} \mathbb{I} \{ \Theta_D \geq \theta_V^2 \} \right) \right\} \\ &\stackrel{\sigma_V^2 \ll \sigma_H^2}{\approx} \frac{2}{3} \phi_{\text{TPLR}}(\mu_V) + \frac{1}{6} \phi_{\text{TPLR}}(\mu_V + \sqrt{3}\sigma_V) + \frac{1}{6} \phi_{\text{TPLR}}(\mu_V - \sqrt{3}\sigma_V), \end{aligned} \quad (6)$$

where

$$\phi_{\text{TPLR}}(x) = 1 - Q_{\frac{1}{2}} \left(\sqrt{\lambda}, \sqrt{\frac{\Theta_D - x^2}{\sigma_H^2}} \mathbb{I} \{ \Theta_D \geq x^2 \} \right). \quad (7)$$

This robust approximation for $\sigma_V^2 \ll \sigma_H^2$ was proposed by [44]–[46]. This robust result becomes more tighter along with the ratio of σ_V^2 / σ_H^2 approaching to zero.

A. Asymptotic Result for TPLR as $\Theta_D \rightarrow 0$

Before presenting the asymptotic analysis for TPLR, we first give the following proposition.

Proposition 1. *The asymptotic expression for $Q_M(a, b)$, the generalized Marcum Q-function, as $b \rightarrow 0$, is given by*

$$Q_M(a, b) \stackrel{b \rightarrow 0}{\simeq} 1 - \frac{\exp\left(-\frac{a^2}{2}\right)}{\Gamma(M+1) 2^M} b^{2M} + o(b^{2M+1}), \quad (8)$$

where M, a, b are non-negative, $o(\cdot)$ and $\Gamma(\cdot)$ denote the higher order term and Gamma function [42], respectively.

Proof. See Appendix B. □

We present some numerical results in Fig. 2 to validate the correctness of the derived asymptotic expression for the generalized Marcum Q-function. It is obvious that the asymptotic results match the exact results very well when b is sufficiently small.

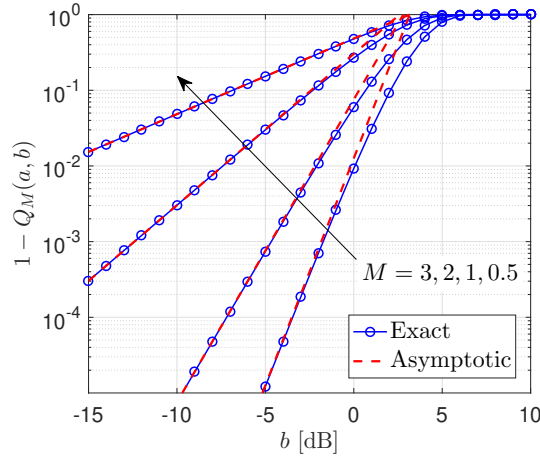


Fig. 2: $1 - Q_M(a, b)$ versus b for $a = 1$.

For $G_D \rightarrow +\infty$, we have

$$\Theta_D = \frac{-1}{G_D} \ln \frac{\lambda_D}{K_1 G_D} \rightarrow 0^+, \quad \text{as } G_D \rightarrow +\infty. \quad (9)$$

In the following, we will analyze the asymptotic behaviour of TPLR when $\Theta_D \rightarrow 0^+$ (or equivalently $G_D \rightarrow +\infty$).

Lemma 1. *For $\Theta_D \rightarrow 0$ (or equivalently $G_D \rightarrow +\infty$), by using Proposition 1, the asymptotic expression for TPLR can be derived as*

$$\text{TPLR} \stackrel{\Theta_D \rightarrow 0}{\simeq} \frac{\sqrt{\pi}}{4\Gamma(1.5)\sigma_V\sigma_H} \exp\left(-\frac{\lambda}{2} - \frac{\mu_V^2}{2\sigma_V^2}\right) \Theta_D. \quad (10)$$

Proof. See Appendix C. □

Remark 2. In the G_D saturation case, a linear mapping from Θ_D to the TPLR is derived, shown in Lemma 1, which is interesting in the performance analysis aspect. This asymptotic expression can not only simplify the TPLR calculation significantly, but also reveal the relationship between the TPLR and Θ_D in the G_D saturation case.

Remark 3. In fact, the asymptotic expression presented in Lemma 1 can be also viewed as the asymptotic result for the outage probability over Beckmann fading channels (including Rayleigh, Hoyt and Rice fading channels), where Θ_D represents the received SNR threshold, and the instantaneous SNR at the receiver is $\gamma = \theta_V^2 + \theta_H^2$.

Although the atmospheric turbulence is not a main investigation in this paper (we will consider this issue in our future work), we can simply analyze this impact on the TPLR in the G_D saturation case based on Lemma 1, shown in Corollary 1 where the turbulence is modeled by a Gamma-Gamma distribution, a widely adopted turbulence model [34].

Corollary 1. In the G_D saturation case, if the atmospheric turbulence modeled by a Gamma-Gamma distribution is considered over the FSO link, the asymptotic result for the TPLR is

$$\begin{aligned} \text{TPLR} = & \frac{\sqrt{\pi}}{4\Gamma(1.5)\sigma_V\sigma_H} \exp\left(-\frac{\lambda}{2} - \frac{\mu_V^2}{2\sigma_V^2}\right) \Theta_D \\ & + \underbrace{\frac{\sqrt{\pi}(\psi(\alpha_D) + \psi(\beta_D) - \ln(\alpha_D\beta_D))}{4G_D\Gamma(1.5)\sigma_V\sigma_H}}_{\text{Atmospheric Turbulence}} \exp\left(-\frac{\lambda}{2} - \frac{\mu_V^2}{2\sigma_V^2}\right), \end{aligned} \quad (11)$$

where $\psi(\cdot)$ denotes the digamma function [42], α_D and β_D are the fading parameters of large-scale and small-scale fluctuations, respectively.

Proof. It is obvious that the TPLR in (2) becomes

$$\begin{aligned} \text{TPLR} &= \Pr\left\{I_D K_1 G_D L(\theta) \geq \lambda_D\right\} = \Pr\left\{\theta^2 \leq \frac{-1}{G_D} \ln \frac{\lambda_D}{I_D K_1 G_D}\right\} \\ &= \Pr\left\{\theta^2 \leq \Theta_D + \frac{1}{G_D} \ln I_D\right\}, \end{aligned} \quad (12)$$

where I_D represents the atmospheric turbulence following a Gamma-Gamma distribution with the PDF [34, Eq. (11)],

$$f_{I_D}(x) = \frac{2(\alpha_D\beta_D)^{\frac{\alpha_D+\beta_D}{2}}}{\Gamma(\alpha_D)\Gamma(\beta_D)} x^{\frac{\alpha_D+\beta_D}{2}-1} K_{\alpha_D-\beta_D}\left(2\sqrt{\alpha_D\beta_D x}\right), \quad (13)$$

where $K(\cdot)$ denotes the modified Bessel function of the second kind [42].

Define $\Theta'_D = \Theta_D + \frac{1}{G_D} \ln I_D$. In view of Lemma 1, it is easy to derive the asymptotic result for the TPLR as

$$\begin{aligned} \text{TPLR} &\stackrel{\Theta_D \rightarrow 0}{\simeq} \frac{\sqrt{\pi}}{4\Gamma(1.5)\sigma_V\sigma_H} \exp\left(-\frac{\lambda}{2} - \frac{\mu_V^2}{2\sigma_V^2}\right) \cdot \mathbb{E}_{I_D}\{\Theta'_D\} \\ &= \frac{\sqrt{\pi}}{4\Gamma(1.5)\sigma_V\sigma_H} \exp\left(-\frac{\lambda}{2} - \frac{\mu_V^2}{2\sigma_V^2}\right) \Theta_D + \underbrace{\frac{\sqrt{\pi}}{4G_D\Gamma(1.5)\sigma_V\sigma_H} \exp\left(-\frac{\lambda}{2} - \frac{\mu_V^2}{2\sigma_V^2}\right) \cdot \mathbb{E}_{I_D}\{\ln I_D\}}_{\text{Atmospheric Turbulence}}. \end{aligned} \quad (14)$$

By using the PDF of I_D , the expectation of $\ln I_D$ with respect to I_D can be obtained as [34]

$$\begin{aligned} \mathbb{E}\{\ln I_D\} &= \frac{2(\alpha_D\beta_D)^{\frac{\alpha_D+\beta_D}{2}}}{\Gamma(\alpha_D)\Gamma(\beta_D)} \times \int_0^\infty \ln(x) \cdot x^{\frac{\alpha_D+\beta_D}{2}-1} K_{\alpha_D-\beta_D}\left(2\sqrt{\alpha_D\beta_D x}\right) dx \\ &= \psi(\alpha_D) + \psi(\beta_D) - \ln(\alpha_D\beta_D), \end{aligned} \quad (15)$$

Combining (14) and (15) yields Corollary 1. \square

Remark 4. Considering $\psi(x) - \ln x \leq 0$ for $x > 0$ in Corollary 1, we can conclude that the impact of atmospheric turbulence is always negative to TPLR in the G_D saturation case. Further, the negative impact is quantified by the second part in (11), compared to the TPLR without atmospheric turbulence in Lemma 1.

B. Special Case for $\mu_V = \mu_H = 0$ and $\sigma_V \neq \sigma_H$

To get the exact closed-form expression for TPLR, we relax the conditions for the statistical characteristics of θ_V and θ_H , i.e., $\mu_V = \mu_H = 0$ and $\sigma_V \neq \sigma_H$. In this simplified case, θ^2 follows the Hoyt distribution, and the corresponding TPLR is given by [43]

$$\text{TPLR} = \frac{2q}{1+q^2} Ie\left(\frac{1-q^2}{1+q^2}, \frac{(1+q^2)^2}{4q^2(\sigma_V^2 + \sigma_H^2)} \Theta_D\right), \quad (16)$$

where $Ie(\cdot, \cdot)$ denotes the Rice Ie -function defined in [43, Eq. (3)], and $q \in [0, 1]$ is given by

$$q = \begin{cases} \sigma_V/\sigma_H, & \text{for } \sigma_V \leq \sigma_H; \\ \sigma_H/\sigma_V, & \text{for } \sigma_V > \sigma_H. \end{cases} \quad (17)$$

From the asymptotic analysis for the general parameter settings, the asymptotic expression in the Hoyt distribution case ($\mu_V = \mu_H = 0$) can be easily derived as

$$\text{TPLR} \stackrel{\Theta_D \rightarrow 0}{\simeq} \frac{\sqrt{\pi}}{4\sigma_V\sigma_H\Gamma(1.5)} \Theta_D. \quad (18)$$

The asymptotic expression can be also easily derived by using the asymptotic result for the generalized Marcum Q-function in Proposition 1, because the Rice I_e -function can be written in the Marcum Q-function form, given by [43]

$$\begin{aligned} I_e(k, x) &= \int_0^x \exp(-t) I_0(kt) dt \\ &= \frac{1}{\sqrt{1-k^2}} \left[Q_1 \left(\sqrt{(1+\sqrt{1-k^2})} x, \sqrt{(1-\sqrt{1-k^2})} x \right) \right. \\ &\quad \left. - Q_1 \left(\sqrt{(1-\sqrt{1-k^2})} x, \sqrt{(1+\sqrt{1-k^2})} x \right) \right]. \end{aligned} \quad (19)$$

When $q = 1$, i.e., $\sigma_V = \sigma_H = \sigma$, the distribution of θ is reduced into the Rayleigh distribution, and the corresponding TPLR becomes

$$\text{TPLR} = 1 - \exp\left(-\frac{\Theta_D}{2\sigma^2}\right). \quad (20)$$

C. Special Case for $\mu_V \neq \mu_H \neq 0$ and $\sigma_V = \sigma_H = \sigma$

For $\mu_V \neq \mu_H$ and $\sigma_V = \sigma_H = \sigma$, i.e., the Rice case, we can rewrite the TPLR as

$$\begin{aligned} \text{TPLR} &= \Pr\{\theta_V^2 + \theta_H^2 \leq \Theta_D\} \\ &= \Pr\left\{ \left(\frac{\theta_V}{\sqrt{\mu_V^2 + \mu_H^2}} \right)^2 + \left(\frac{\theta_H}{\sqrt{\mu_V^2 + \mu_H^2}} \right)^2 \leq \frac{\Theta_D}{\mu_V^2 + \mu_H^2} \right\} \\ &= \Pr\left\{ \sqrt{\left(\frac{\theta_V}{\sqrt{\mu_V^2 + \mu_H^2}} \right)^2 + \left(\frac{\theta_H}{\sqrt{\mu_V^2 + \mu_H^2}} \right)^2} \leq \sqrt{\frac{\Theta_D}{\mu_V^2 + \mu_H^2}} \right\}. \end{aligned} \quad (21)$$

Let $\sin \beta = \frac{\mu_V}{\sqrt{\mu_V^2 + \mu_H^2}}$ and $\cos \beta = \frac{\mu_H}{\sqrt{\mu_V^2 + \mu_H^2}}$. It is easy to see that

$$\begin{aligned} \frac{\theta_V}{\sqrt{\mu_V^2 + \mu_H^2}} &\sim \mathcal{N}\left(\sin \beta, \frac{\sigma^2}{\mu_V^2 + \mu_H^2}\right), \\ \frac{\theta_H}{\sqrt{\mu_V^2 + \mu_H^2}} &\sim \mathcal{N}\left(\cos \beta, \frac{\sigma^2}{\mu_V^2 + \mu_H^2}\right). \end{aligned}$$

From the definition of the Rice distribution, we can conclude that

$$\sqrt{\left(\frac{\theta_V}{\sqrt{\mu_V^2 + \mu_H^2}} \right)^2 + \left(\frac{\theta_H}{\sqrt{\mu_V^2 + \mu_H^2}} \right)^2} \sim \text{Rice}\left(1, \frac{\sigma}{\sqrt{\mu_V^2 + \mu_H^2}}\right),$$

and the corresponding TPLR can be easily derived by using the well-known standard cumulative distribution function (CDF) of the Rice distribution, given by

$$\text{TPLR} = 1 - Q_1\left(\frac{\sqrt{\mu_V^2 + \mu_H^2}}{\sigma}, \frac{\sqrt{\Theta_D}}{\sigma}\right). \quad (22)$$

From the derivation in Proposition 1, when $\Theta_D \rightarrow 0$, the asymptotic TPLR is

$$\text{TPLR} \stackrel{\Theta_D \rightarrow 0}{\simeq} \exp\left(-\frac{\mu_V^2 + \mu_H^2}{2\sigma^2}\right) \frac{1}{2\sigma^2} \Theta_D + o(\Theta_D^2). \quad (23)$$

IV. TRANSMISSION PROBABILITY DEPENDING ONLY ON EAVESDROPPER

The transmission probability depending only on eavesdropper (TPE) is defined as the probability that the received power at the eavesdropper is less than a threshold (λ_E), i.e.,

$$\begin{aligned} \text{TPE} &= \Pr\left\{K_2 P_D(\theta) L(\theta_E) G_D \leq \lambda_E\right\} = \Pr\left\{K_1 K_2 G_D^2 \exp(-G_D(2\theta^2 + 2\alpha\theta_V + \alpha^2)) \leq \lambda_E\right\} \\ &= \Pr\left\{\theta^2 + \alpha\theta_V \geq \underbrace{\frac{-1}{2G_D} \ln \frac{\lambda_E}{K_1 K_2 G_D^2} - \frac{\alpha^2}{2}}_{\Theta_E}\right\}. \end{aligned} \quad (24)$$

An approximation for the TPE will be given in Theorem 2 based on the GHQ.

Theorem 2. *The TPE can be approximated as*

$$\text{TPE} \approx \frac{1}{\sqrt{\pi}} \sum_{i=1}^N \omega_i f_{\text{TPE}}(x_i), \quad (25)$$

where N , ω_i , and x_i are the same as those in (4), and $f_{\text{TPE}}(x)$ is given in (26),

$$\begin{aligned} f_{\text{TPE}}(x) &= Q_{\frac{1}{2}}\left(\sqrt{\lambda}, \sqrt{\frac{\Theta_E - (\sqrt{2}\sigma_V x + \mu_V)^2 - \alpha(\sqrt{2}\sigma_V x + \mu_V)}{\sigma_H^2}}\right) \\ &\quad \times \mathbb{I}\left\{\Theta_E \geq (\sqrt{2}\sigma_V x + \mu_V)^2 + \alpha(\sqrt{2}\sigma_V x + \mu_V)\right\}. \end{aligned} \quad (26)$$

Proof. See Appendix D. □

From [44]–[46], for $\sigma_V^2 \ll \sigma_H^2$, the TPE can be robustly approximated by

$$\text{TPE} \stackrel{(a)}{\approx} \frac{2}{3} \phi_{\text{TPE}}(\mu_V) + \frac{1}{6} \phi_{\text{TPE}}(\mu_V + \sqrt{3}\sigma_V) + \frac{1}{6} \phi_{\text{TPE}}(\mu_V - \sqrt{3}\sigma_V), \quad (27)$$

where (a) follows [44, Eq. (4)], and $\phi_{\text{TPE}}(\cdot)$ is given by

$$\phi_{\text{TPE}}(x) = Q_{\frac{1}{2}}\left(\sqrt{\lambda}, \sqrt{\frac{\Theta_E - x^2 - \alpha x}{\sigma_H^2}} \mathbb{I}\{\Theta_E \geq x^2 + \alpha x\}\right). \quad (28)$$

A. Analysis on Θ_E versus G_D

Θ_E is given by

$$\Theta_E = \frac{-1}{2G_D} \ln \frac{\lambda_E}{K_1 K_2 G_D^2} - \frac{\alpha^2}{2}. \quad (29)$$

The first derivative of Θ_E with respect to G_D is

$$\frac{\partial \Theta_E}{\partial G_D} = \frac{\ln \frac{\lambda_E}{K_1 K_2 G_D^2} + 2}{2G_D^2}. \quad (30)$$

Let $\frac{\partial \Theta_E}{\partial G_D} = 0$. The root for $G_D > 0$ is

$$G_D^* = e \sqrt{\frac{\lambda_E}{K_1 K_2}}. \quad (31)$$

We can conclude that Θ_E is an increasing function over $G_D \in (0, G_D^*)$, and a decreasing function over $G_D \in (G_D^*, \infty)$. The maximum value of Θ_E is

$$\Theta_E^* = \Theta_E|_{G_D=G_D^*} = \frac{\sqrt{K_1 K_2}}{e \sqrt{\lambda_E}} - \frac{\alpha^2}{2}. \quad (32)$$

For $G_D \rightarrow \infty$, we have

$$\lim_{G_D \rightarrow \infty} \frac{-1}{2G_D} \ln \frac{\lambda_E}{K_1 K_2 G_D^2} - \frac{\alpha^2}{2} = -\frac{\alpha^2}{2}. \quad (33)$$

For $G_D \rightarrow 0^+$, we have the limit

$$\lim_{G_D \rightarrow 0^+} \frac{-1}{2G_D} \ln \frac{\lambda_E}{K_1 K_2 G_D^2} - \frac{\alpha^2}{2} \rightarrow -\infty. \quad (34)$$

Remark 5. Combining the analysis on Θ_E and (24), we can conclude that the TPE is decreasing over $G_D \in (0, G_D^*)$ due to the increase in Θ_E , while the TPE is increasing over $G_D \in (G_D^*, \infty)$ due to the decrease in Θ_E . When G_D is sufficiently large, the TPE will increase to an upper bound, because Θ_E goes to a constant $(-\alpha^2/2)$. If $G_D \rightarrow 0^+$, $\Theta_E \rightarrow -\infty$, resulting in the TPE going to 1.

B. Asymptotic Analysis for TPE

Lemma 2. When $\Theta_E \rightarrow \frac{-\alpha^2}{4}$ from the right side in the real number axis, the asymptotic result for TPE is given by

$$\text{TPE} \stackrel{\Theta_E \rightarrow \frac{-\alpha^2}{4}}{\simeq} 1 - \frac{\sqrt{\pi} (\alpha^2 + 4\Theta_E)}{16\sigma_H \sigma_V \Gamma(1.5)} \exp\left(-\frac{\lambda}{2} - \frac{(\alpha/2 + \mu_V)^2}{2\sigma_V^2}\right) \times \mathbb{I}\left\{\Theta_E \geq \frac{-\alpha^2}{4}\right\}. \quad (35)$$

For $\Theta_E \leq \frac{-\alpha^2}{4}$, the TPE is exactly equal to 1.

Proof. See Appendix E. □

Remark 6. α reflects the distance between the sender and the eavesdropper. Specifically, $\alpha \rightarrow 0$ means that the eavesdropper is close to the sender. Lemma 2 presents a quantitative relationship between Θ_E (or G_D) and α . When $\Theta_E \leq \frac{-\alpha^2}{4}$, the event of received power at the eavesdropper below the pre-set threshold (λ_E) happens with probability one. Further, Lemma 2 reveals the linear trend of TPE with respect to Θ_E when $\Theta_E \rightarrow \left(\frac{-\alpha^2}{4}\right)^+$.

V. TRANSMISSION PROBABILITY DEPENDING ON BOTH LEGITIMATE RECEIVER AND EAVESDROPPER

The transmission probability depending on both the legitimate receiver and the eavesdropper (TPRE) is defined as

$$\text{TPRE} = \Pr \left\{ P_D(\theta) \geq \lambda_D, P_E(\theta_E) \leq \lambda_E \right\} = \Pr \left\{ K_1 G_D L(\theta) \geq \lambda_D, K_1 K_2 G_D^2 L(\theta) L(\theta_E) \leq \lambda_E \right\}, \quad (36)$$

which can be further written by using the definition of Θ_D and Θ_E ,

$$\text{TPRE} = \Pr \left\{ \theta_H^2 \leq \Theta_D - \theta_V^2, \theta_H^2 \geq \Theta_E - \theta_V^2 - \alpha \theta_V \right\} = \Pr \left\{ \frac{\Theta_E - \theta_V^2 - \alpha \theta_V}{\sigma_H^2} \leq \frac{\theta_H^2}{\sigma_H^2} \leq \frac{\Theta_D - \theta_V^2}{\sigma_H^2} \right\}. \quad (37)$$

By applying the GHQ method, an approximate result for the TPRE is derived in Theorem 3.

Theorem 3. *A closed-form expression for an approximate TPRE is*

$$\text{TPRE} \approx \frac{1}{\sqrt{\pi}} \sum_{i=1}^N \omega_i f_{\text{TPRE}}(x_i), \quad (38)$$

where N , ω_i , x_i are the same as those in (4), and $f_{\text{TPRE}}(\cdot)$ is given in (39),

$$\begin{aligned} f_{\text{TPRE}}(x) = \mathbb{I} \left\{ x \geq \frac{\Theta_E - \Theta_D}{\alpha \sqrt{2} \sigma_V} - \frac{\mu_V}{\sqrt{2} \sigma_V} \right\} & \left[Q_{\frac{1}{2}} \left(\sqrt{\lambda}, \frac{1}{\sigma_H} \sqrt{\Theta_E - (\sqrt{2} \sigma_V x + \mu_V)^2 - \alpha (\sqrt{2} \sigma_V x + \mu_V)} \right) \right. \\ & \times \mathbb{I} \left\{ \Theta_E \geq (\sqrt{2} \sigma_V x + \mu_V)^2 + \alpha (\sqrt{2} \sigma_V x + \mu_V) \right\} \left. \right) \\ & - Q_{\frac{1}{2}} \left(\sqrt{\lambda}, \frac{1}{\sigma_H} \sqrt{\left(\Theta_D - (\sqrt{2} \sigma_V x + \mu_V)^2 \right) \mathbb{I} \left\{ \Theta_D \geq (\sqrt{2} \sigma_V x + \mu_V)^2 \right\}} \right) \left. \right]. \quad (39) \end{aligned}$$

Proof. See Appendix F. □

For $\sigma_V^2 \ll \sigma_H^2$, according to [44]–[46], the TPRE in (73) can be robustly approximated by

$$\text{TPRE} \stackrel{\sigma_V^2 \ll \sigma_H^2}{\approx} \frac{2}{3} \phi_{\text{TPRE}}(\mu_V) + \frac{1}{6} \phi_{\text{TPRE}}(\mu_V + \sqrt{3} \sigma_V) + \frac{1}{6} \phi_{\text{TPRE}}(\mu_V - \sqrt{3} \sigma_V), \quad (40)$$

where $\phi_{\text{TPRE}}(x)$ is given in (41),

$$\begin{aligned} \phi_{\text{TPRE}}(x) = \mathbb{I} \left\{ x \geq \frac{\Theta_E}{2\alpha} - \frac{\Theta_D}{\alpha} \right\} & \left[Q_{\frac{1}{2}} \left(\sqrt{\lambda}, \sqrt{\left(\frac{\Theta_E}{2\sigma_H^2} - \frac{x^2 - \alpha x}{\sigma_H^2} \right) \mathbb{I} \left\{ \Theta_E \geq 2(x^2 + \alpha x) \right\}} \right) \right. \\ & \left. - Q_{\frac{1}{2}} \left(\sqrt{\lambda}, \sqrt{\frac{\Theta_D - x^2}{\sigma_H^2} \mathbb{I} \left\{ \Theta_D \geq x^2 \right\}} \right) \right]. \quad (41) \end{aligned}$$

A. Asymptotic Analysis

Lemma 3. *The asymptotic expression for TPRE as $G_D \rightarrow +\infty$ is the same as that for TPLR, given by*

$$\text{TPRE} \stackrel{G_D \rightarrow +\infty}{\simeq} \frac{\sqrt{\pi}\Theta_D}{4\sigma_H\sigma_V\Gamma(1.5)} \exp\left(-\frac{\lambda}{2} - \frac{\mu_V^2}{2\sigma_V^2}\right). \quad (42)$$

Proof. The proof by solid mathematical manipulations is removed due to the page limitation. \square

Lemma 3 can be explained by the fact that for $G_D \rightarrow +\infty$, $\Theta_E \rightarrow \frac{-\alpha^2}{2}$, and $\Theta_D \rightarrow 0^+$. From the previous analysis, we know that for $\Theta_E < \frac{-\alpha^2}{4}$, there is no real root for $\Theta_E - \theta_V^2 - \alpha\theta_V = 0$, i.e., $\Theta_E - \theta_V^2 - \alpha\theta_V$ is always less than zero for any θ_V . From the second equal sign in (37), for $\Theta_E < \frac{-\alpha^2}{4}$, the TPRE becomes $\text{TPRE} = \Pr\left\{\frac{\theta_H^2}{\sigma_H^2} \leq \frac{\Theta_D - \theta_V^2}{\sigma_H^2}\right\}$, which is exactly the definition of TPLR.

Remark 7. *The asymptotic result for the TPRE in Lemma 3 reveals that although the increase in G_D also induces an increase in the received power at the eavesdropper, the positive impact of increasing G_D on the TPRE will domain the overall performance when $G_D \rightarrow \infty$, and more exactly, there is a convergence trend of TPLR and TPRE (i.e., the eavesdropping impact vanishes) in the G_D saturation case.*

B. Special Case for $\mu_V = \mu_H = 0$ and $\sigma_V = \sigma_H = \sigma$

As the TPRE cannot be solved in an exact closed-form, even for the simplest case, i.e., Rayleigh distribution. Here, we only analyze the special case for the TPRE in the Rayleigh case ($\mu_V = \mu_H = 0$ and $\sigma_V = \sigma_H = \sigma$).

Let $X = \theta_V^2 + \theta_H^2$ and $Y = 2\theta_V^2 + 2\theta_H^2 + 2\alpha\theta_V = 2X + 2\alpha\theta_V$. The TPRE can be rewritten as

$$\begin{aligned} \text{TPRE} &= \Pr\left\{K_1 G_D \exp(-G_D X) \geq \lambda_D, \quad K_1 K_2 G_D^2 \exp(-G_D(Y + \alpha^2)) \leq \lambda_E\right\} \\ &= \Pr\left\{X \leq \frac{-1}{G_D} \ln \frac{\lambda_D}{K_1 G_D}, \quad Y \geq \frac{-1}{G_D} \ln \frac{\lambda_E \exp(G_D \alpha^2)}{K_1 K_2 G_D^2}\right\}. \end{aligned} \quad (43)$$

Let

$$\Theta_D = \frac{-1}{G_D} \ln \frac{\lambda_D}{K_1 G_D}, \quad \Theta_E = \frac{-1}{G_D} \ln \frac{\lambda_E \exp(G_D \alpha^2)}{K_1 K_2 G_D^2}.$$

The TPRE is further written as

$$\Pr\{X \leq \Theta_D, Y \geq \Theta_E\} = \int_0^{\Theta_D} \int_{\Theta_E}^{\infty} f_{X,Y}(x, y) dy dx, \quad (44)$$

where $f_{X,Y}(\cdot, \cdot)$ represents the joint PDF of X and Y . From the derivation of $f_{X,Y}(\cdot, \cdot)$ in Appendix G, we can write the TPRE in an integral form as

$$\text{TPRE} = \frac{1}{4\pi\sigma^2\alpha} \int_0^{\Theta_D} \exp\left(-\frac{x}{2\sigma^2}\right) \int_{\max\{\Theta_E, -2\alpha\sqrt{x}+2x\}}^{\infty} \left[x - \left(\frac{y-2x}{2\alpha}\right)^2\right]^{-\frac{1}{2}} dy dx. \quad (45)$$

By considering the derivation in Appendix H, the TPRE can be easily approximated by

$$\text{TPRE} \approx \frac{\Theta_D}{8\pi\alpha\sigma^2} \sum_{i=0}^N \omega'_i f'_{\text{TPRE}} \left(\frac{\Theta_D}{2} x'_i + \frac{\Theta_D}{2} \right), \quad (46)$$

where ω'_i and x'_i are the weights and selected points over the standard integral interval (i.e., [-1,1]) in the Gaussian quadrature respectively, and $f'_{\text{TPRE}}(x)$ is given by

$$f'_{\text{TPRE}}(x) = \exp\left(-\frac{x}{2\sigma^2}\right) \left[\alpha\pi - 2\alpha \arcsin\left(\frac{\max\{\Theta_E, -2\alpha\sqrt{x} + 2x\} - 2x}{2\alpha\sqrt{x}}\right) \right]. \quad (47)$$

Remark 8. *There is no special function involved in (46), resulting in a much faster calculation for TPRE. Moreover, the built-in function for calculating the non-integer order generalized Marcum Q-function is not available in some softwares, such as Matlab, which brings more difficulty to implement Theorem 3.*

C. Further Simplification for the Rayleigh Case

If we let $\max\{\Theta_E, -2\alpha\sqrt{x} + 2x\}$ be always $-2\alpha\sqrt{x} + 2x$. The integral form of the TPRE in (45) becomes

$$\text{TPRE} = \frac{1}{4\pi\sigma^2\alpha} \int_0^{\Theta_D} \exp\left(-\frac{x}{2\sigma^2}\right) \int_{-2\alpha\sqrt{x}+2x}^{2x+2\alpha\sqrt{x}} \left[x - \left(\frac{y-2x}{2\alpha}\right)^2 \right]^{-\frac{1}{2}} dy dx. \quad (48)$$

By using the following integral identity,

$$\int_{-2\alpha\sqrt{x}+2x}^{2x+2\alpha\sqrt{x}} \left[x - \left(\frac{y-2x}{2\alpha}\right)^2 \right]^{-\frac{1}{2}} dy = 2\alpha\pi, \quad (49)$$

the closed-form expression for the TPRE can be finally derived as

$$\text{TPRE} = 1 - \exp\left(-\frac{\Theta_D}{2\sigma^2}\right). \quad (50)$$

If $G_D \rightarrow +\infty$, we have $\Theta_D = \frac{1}{G_D} \ln \frac{K_1 G_D}{\lambda_D} \rightarrow 0^+$. In this case ($\Theta_D \rightarrow 0^+$), by using $\exp(-x) \simeq 1-x$ for $x \rightarrow 0$, the asymptotic result for TPRE is

$$\text{TPRE} \simeq \frac{\Theta_D}{2\sigma^2} = \frac{1}{2\sigma^2 G_D} \ln \frac{K_1 G_D}{\lambda_D}. \quad (51)$$

From the previous analysis, we know that the condition for using this very simple expression (50) is $\Theta_E < \frac{-\alpha^2}{4}$. In this condition, the impact on the TPRE from the eavesdropper vanishes.

VI. NUMERICAL RESULTS

A. TPLR Simulations

In this subsection, we run some Monte-Carlo simulations to validate the derived closed-form expressions for the TPLR. By referring to the parameter settings in [31], [36], some selected simulation results

are shown in Figs. 3–6, where 10^7 realizations are generated to get each average result according to the statistical properties, i.e., μ_V , μ_H , σ_V and σ_H .

In Fig. 3, the growing trend of TPLR with decreasing σ_H^2 is obvious due to more weaker vibrations. As expected, the TPLR is decreasing as Θ_D approaches to zero (or equivalently $G_D \rightarrow +\infty$).

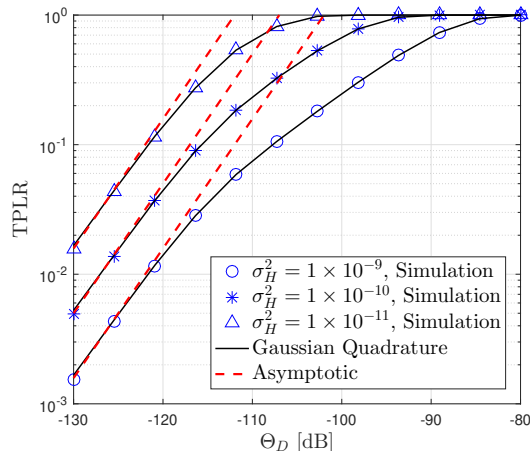


Fig. 3: TPLR versus Θ_D for $\mu_V = 1 \times 10^{-8}$, $\mu_H = 5 \times 10^{-8}$, $\sigma_V^2 = 1 \times 10^{-12}$ and $N = 300$ for Gaussian quadrature.

Figs. 4–5 plot the TPLR versus Θ_D for different variance cases in the Hoyt and Rice models respectively, where the trends with respect to Θ_D (or variance) are the same as those in Fig. 3. The proposed asymptotic results match the exact results very well especially in the low Θ_D region in Figs. 3–5. It is worth noting that the asymptotic results are calculated much faster than the results by the Gaussian quadrature rule, which provides an alternative and efficient method, especially when we need a very high accuracy (several thousand terms may be needed in Gaussian quadrature method).

The results shown in Fig. 6 validate the high accuracy of the robust approximation proposed in [44]–[46], especially for $\sigma_V^2 \ll \sigma_H^2$. We can easily see that the gap between the exact and robust results almost vanishes as the ratio of σ_V^2/σ_H^2 approaches to zero.

B. TPE Simulations

In this subsection, we run the Monte-Carlo simulation to validate the correctness of the derived closed-form expressions for the TPE. To facilitate the simulation setting, we assume that $P_S = 0$ dB, $G_S = G_E = 10^9$, $\eta_S = \eta_D = \eta_E = 0.9$, $\eta_q = 0.1$, $\eta_B = 0.04$, $\lambda_1 = \lambda_2 = 780$ nm, $Z_1 = Z_2 = 900$ km, and $L_A(D_1) = L_A(D_2) = 0.5$, by considering Table I in [31]. As the robust approximation for the

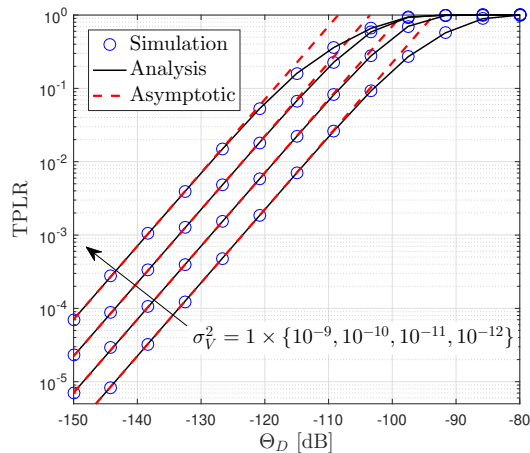


Fig. 4: TPLR in the Hoyt case versus Θ_D for $\mu_V = \mu_H = 0$, and $\sigma_H^2 = 5 \times 10^{-11}$.

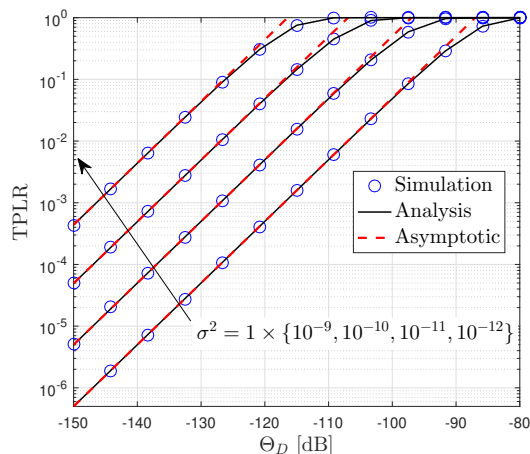


Fig. 5: TPLR in the Rice case versus Θ_D for $\sigma_V = \sigma_H = \sigma$, $\mu_V = 1 \times 10^{-7}$ and $\mu_H = 5 \times 10^{-7}$.

Gaussian distribution has been investigated very well in [44]–[46], we do not validate the high accuracy for $\sigma_V^2 \ll \sigma_H^2$ in the following simulations.

As shown in Fig. 7, the TPE remains 1 before a sharp decrease up to the lowest bound, and after this lowest bound, the TPE grows rapidly to 1 as G_D increases, which is exactly as the TPE changing analysis in Remark 5. To show the lowest bound, Fig. 8 uses the log-scale to plot TPE in Fig. 7, where this lowest bound grows with increasing μ_H . From Figs. 7–8, it is obvious that a large μ_H results in a large TPE, which can be explained by the fact that the received power performance at the legitimate receiver becomes worse, thereby accordingly decreasing the received power at the eavesdropper.

Fig. 9 plots the TPE versus Θ_E from $-\frac{\alpha^2}{4}$ to $15\alpha^2$. As analyzed in the IV-B subsection, for $\Theta_E \rightarrow -\frac{\alpha^2}{4}$ from the right side in the real number axis, the TPE can be approximated by a linear function. Moreover,

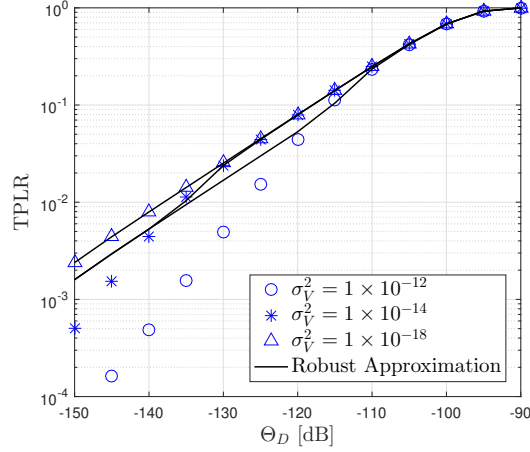


Fig. 6: TPLR versus Θ_D for $\sigma_H^2 = 1 \times 10^{-10}$, $\mu_V = 1 \times 10^{-8}$ and $\mu_H = 5 \times 10^{-8}$.

there is a decreasing trend in the TPE with respect to α .

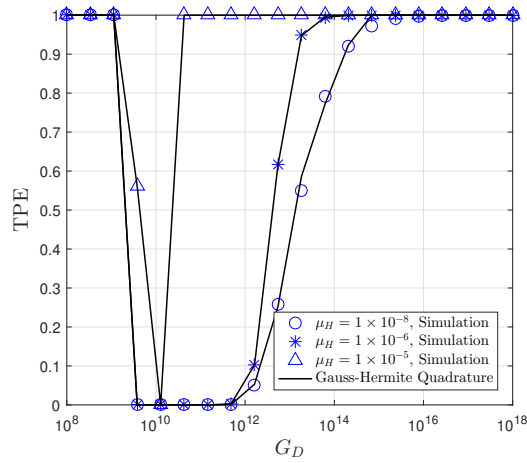


Fig. 7: TPE versus G_D for $\mu_V = 1 \times 10^{-7}$, $\sigma_V^2 = 1 \times 10^{-12}$, $\sigma_H^2 = 1 \times 10^{-13}$, $\lambda_E = 1 \times 10^{-20}$, $\alpha = 1 \times 10^{-9}$, and $N = 30$ for the Gaussian quadrature.

C. TPPE Simulations

In this subsection, the same system parameter settings as those in the first paragraph in the VI-B subsection are assumed for convenience.

In Fig. 10, the TPPE versus G_D is plotted, where a fluctuation is obvious in the medium G_D region, while a monotonic decreasing is presented in the high G_D region, and this decreasing trend can be

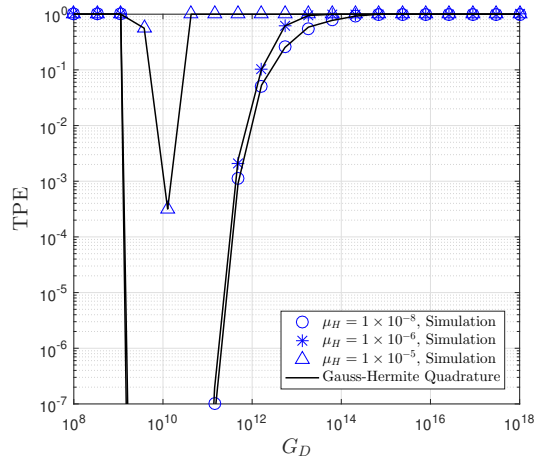


Fig. 8: TPE versus G_D in the log-scale of Fig. 7.

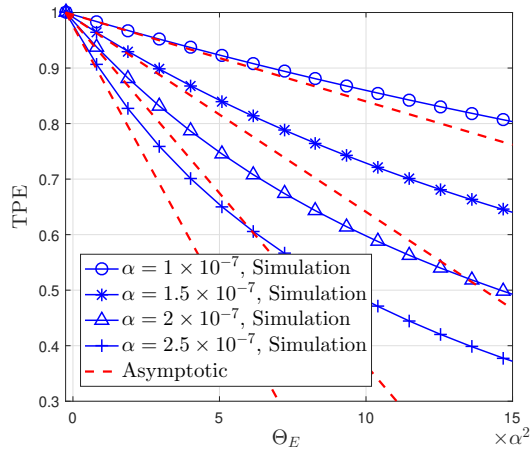


Fig. 9: TPE versus Θ_E for $\mu_V = 1 \times 10^{-7}$, $\mu_H = 1 \times 10^{-8}$, $\sigma_V^2 = 1 \times 10^{-12}$, $\sigma_H^2 = 1 \times 10^{-13}$, and $\lambda_E = 1 \times 10^{-20}$.

approximated by a linear function (independent on α , which has been proved in the asymptotic analysis for TPPE) in the log-scale. A smaller α means a more close eavesdropper around the sender, which results in the decrease in the TPPE.

There also exists a fluctuation in $(\lambda_D, \lambda_E) = (10^{-15}, 10^{-20}), (10^{-20}, 10^{-20})$ cases in Fig. 11. In contrast, the TPPE for $(\lambda_D, \lambda_E) = (10^{-15}, 10^{-15})$ remains constant (equal to 1) before a monotonic decrease. From Fig. 11, we can also observe that the TPPE is an increasing function with respect to λ_E , and a decreasing function with respect to λ_D , which is easily explained by the joint probability properties in (36).

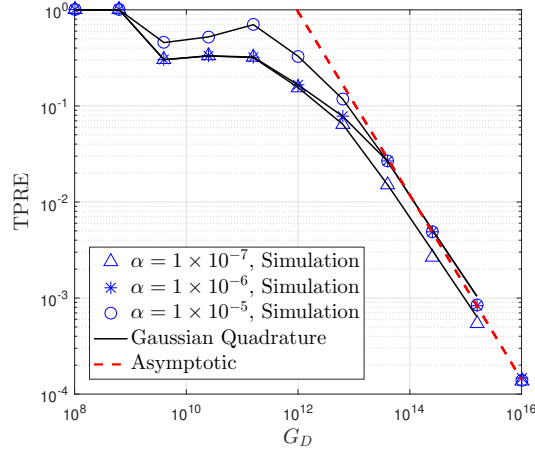


Fig. 10: TPRE versus G_D for $\mu_V = 1 \times 10^{-8}$, $\mu_H = 5 \times 10^{-8}$, $\sigma_V^2 = 1 \times 10^{-12}$, and $\sigma_H^2 = 1 \times 10^{-10}$, $\lambda_D = 1 \times 10^{-15}$, $\lambda_E = 1 \times 10^{-20}$, and $N = 300$ for Gaussian quadrature.

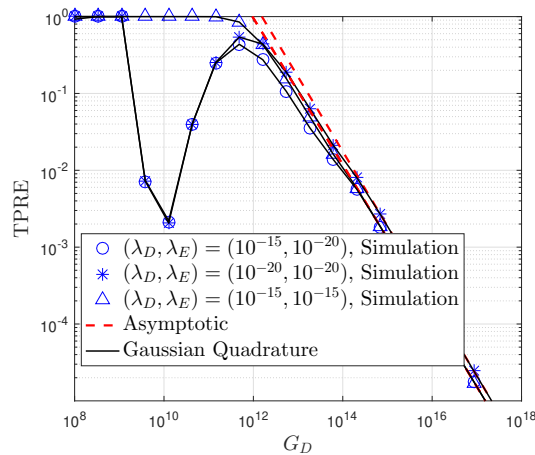


Fig. 11: TPRE versus G_D for $\mu_V = \mu_H = 0$, $\sigma_V^2 = \sigma_H^2 = 1 \times 10^{-11}$ (Rayleigh distribution), $\alpha = 1 \times 10^{-6}$, and $N = 110$.

VII. CONCLUSION

In this paper, closed-form expressions for TPLR, TPE, and TPRE were derived based on the Gaussian quadrature rule, as well as the corresponding asymptotic formulas valid in the high telescope gain at the legitimate receiver region by using the asymptotic result for the generalized Marcum Q-function which could be also employed to derive the asymptotic expression for outage probability (showing the diversity order and array gain) over Beckmann fading channels. The approximate expressions based on the Gaussian quadrature rule needs many terms for a high accuracy and the computation becomes more

slower with increasing the terms. Alternatively, we can use the asymptotic expressions to get results with a very high accuracy in the high region of the telescope gain at the legitimate receiver. Moreover, some closed-form expressions or more concise expressions in some special cases (like Rayleigh, Hoyt and Rice pointing error model cases) were also derived for much faster computations.

APPENDIX A: PROOF OF THEOREM 1

We can rewrite (2) as

$$\text{TPLR} = \Pr \left\{ \theta_V^2 + \theta_H^2 \leq \Theta_D \right\} = \Pr \left\{ \theta_H^2 \leq \Theta_D - \theta_V^2 \right\} = \Pr \left\{ \frac{\theta_H^2}{\sigma_H^2} \leq \frac{\Theta_D - \theta_V^2}{\sigma_H^2} \right\}, \quad (52)$$

where θ_H^2/σ_H^2 follows the non-central Chi-squared distribution with the unit degree, and the corresponding cumulative distribution function (CDF) is given by [41]

$$F_{\theta_H^2/\sigma_H^2}(x) = 1 - Q_{\frac{1}{2}} \left(\sqrt{\lambda}, \sqrt{x \mathbb{I}\{x \geq 0\}} \right). \quad (53)$$

By using the CDF of θ_H^2/σ_H^2 and iterative expectation operation, the TPLR can be derived as

$$\begin{aligned} \text{TPLR} &= \mathbb{E}_{\theta_V} \left\{ 1 - Q_{\frac{1}{2}} \left(\sqrt{\lambda}, \sqrt{\frac{\Theta_D - \theta_V^2}{\sigma_H^2} \mathbb{I}\{\Theta_D \geq \theta_V^2\}} \right) \right\} \\ &\stackrel{(a)}{=} 1 - \frac{1}{\sqrt{2\pi\sigma_V^2}} \int_{-\infty}^{+\infty} \exp \left(-\frac{(\theta_V - \mu_V)^2}{2\sigma_V^2} \right) Q_{\frac{1}{2}} \left(\sqrt{\lambda}, \sqrt{\frac{\Theta_D - \theta_V^2}{\sigma_H^2} \mathbb{I}\{\Theta_D \geq \theta_V^2\}} \right) d\theta_V, \end{aligned} \quad (54)$$

where $\mathbb{E}\{\cdot\}$ denotes the expectation operator, and (a) follows the PDF of θ_V , i.e., PDF of Gaussian distribution. Let $x = \frac{\theta_V - \mu_V}{\sqrt{2}\sigma_V}$ in (54), and then, we can get the integral form as (55),

$$\text{TPLR} = 1 - \frac{1}{\sqrt{\pi}} \int_{-\infty}^{+\infty} \exp(-x^2) Q_{\frac{1}{2}} \left(\sqrt{\lambda}, \sqrt{\frac{\Theta_D - (\sqrt{2}\sigma_V x + \mu_V)^2}{\sigma_H^2} \mathbb{I}\{\Theta_D \geq (\sqrt{2}\sigma_V x + \mu_V)^2\}} \right) dx, \quad (55)$$

which can be approximated by using the Gaussian quadrature rule, shown in Theorem 1.

APPENDIX B: PROOF OF PROPOSITION 1

The Marcum Q-function is defined in the integral form,

$$Q_M(a, b) = \int_b^{\infty} x \left(\frac{x}{a} \right)^{M-1} \exp \left(-\frac{x^2 + a^2}{2} \right) I_{M-1}(ax) dx, \quad (56)$$

where $I(\cdot)$ denotes the modified Bessel function of the first kind [42]. To derive the asymptotic result for the Marcum Q-function, we first express $I(\cdot)$ into the infinite series form,

$$I_{M-1}(ax) = \sum_{n=0}^{\infty} \frac{1}{n! \Gamma(n+M)} \left(\frac{ax}{2} \right)^{2n+M-1}. \quad (57)$$

When $x \rightarrow 0$, it is easy to see that the leading term (i.e., the lowest order term of x) for approximating $I(\cdot)$ in (57) is

$$I_{M-1}(ax) \simeq \frac{1}{n!\Gamma(n+M)} \left(\frac{ax}{2}\right)^{2n+M-1} \Big|_{n=0} = \frac{1}{\Gamma(M)} \left(\frac{ax}{2}\right)^{M-1}. \quad (58)$$

By using the asymptotic result for $I(\cdot)$, the asymptotic result for the Marcum Q-function, as $b \rightarrow 0$, can be derived as

$$\begin{aligned} Q_M(a, b) &= \underbrace{\int_0^\infty x \left(\frac{x}{a}\right)^{M-1} \exp\left(-\frac{x^2+a^2}{2}\right) I_{M-1}(ax) dx}_{=1} \\ &\quad - \int_0^b x \left(\frac{x}{a}\right)^{M-1} \exp\left(-\frac{x^2+a^2}{2}\right) I_{M-1}(ax) dx \\ &\stackrel{b \rightarrow 0}{\simeq} 1 - \frac{\exp\left(-\frac{a^2}{2}\right)}{\Gamma(M) 2^{M-1}} \int_0^b x^{2M-1} \exp\left(-\frac{x^2}{2}\right) dx, \end{aligned} \quad (59)$$

where the integral in the last step can be easily derived in a closed-form by using the definition of the lower incomplete Gamma function [42], given by $\int_0^b x^{2M-1} \exp\left(-\frac{x^2}{2}\right) dx = 2^{M-1} \Upsilon\left(M, \frac{b^2}{2}\right)$, where $\Upsilon(\cdot, \cdot)$ denotes the lower incomplete Gamma function. By using the asymptotic result for $\Upsilon(\cdot, \cdot)$, i.e., $\Upsilon(s, x) \rightarrow \frac{x^s}{s}$ as $x \rightarrow 0$, we can finally derive the asymptotic result for the Marcum Q-function as (8).

APPENDIX C: PROOF OF LEMMA 1

From the derivation in Proposition 1, the CDF of the non-central Chi-squared distribution can be approximated by

$$F_{\theta_H^2/\sigma_H^2}(x) = 1 - Q_{\frac{1}{2}}\left(\sqrt{\lambda}, \sqrt{x} \mathbb{I}\{x \geq 0\}\right) \stackrel{x \rightarrow 0}{\simeq} \frac{\exp\left(-\frac{\lambda}{2}\right)}{\Gamma(1.5) \sqrt{2}} x^{0.5} \mathbb{I}\{x \geq 0\}. \quad (60)$$

Substituting the asymptotic expression for $F_{\theta_H^2/\sigma_H^2}(\cdot)$ into (52) yields

$$\begin{aligned} \text{TPLR} &\simeq \frac{\exp\left(-\frac{\lambda}{2}\right)}{\Gamma(1.5) \sqrt{2}} \mathbb{E}_{\theta_V} \left\{ \left(\frac{\Theta_D - \theta_V^2}{\sigma_H^2}\right)^{\frac{1}{2}} \mathbb{I}\{\Theta_D \geq \theta_V^2\} \right\} \\ &= \frac{\exp\left(-\frac{\lambda}{2}\right)}{\Gamma(1.5) \sqrt{2}} \frac{1}{\sqrt{2\pi\sigma_V^2}} \int_{-\infty}^{+\infty} \exp\left(-\frac{(\theta_V - \mu_V)^2}{2\sigma_V^2}\right) \left(\frac{\Theta_D - \theta_V^2}{\sigma_H^2}\right)^{\frac{1}{2}} \mathbb{I}\{\Theta_D \geq \theta_V^2\} d\theta_V \\ &= \frac{\exp\left(-\frac{\lambda}{2}\right)}{\Gamma(1.5) \sqrt{2}} \frac{1}{\sqrt{2\pi\sigma_V^2}} \int_{-\sqrt{\Theta_D}}^{\sqrt{\Theta_D}} \exp\left(-\frac{(\theta_V - \mu_V)^2}{2\sigma_V^2}\right) \left(\frac{\Theta_D - \theta_V^2}{\sigma_H^2}\right)^{\frac{1}{2}} d\theta_V. \end{aligned} \quad (61)$$

As $\Theta_D \rightarrow 0$, the exponential function in the last step in (61) can be approximated by

$$\exp\left(-\frac{(\theta_V - \mu_V)^2}{2\sigma_V^2}\right) \stackrel{\theta_V \rightarrow 0}{\simeq} \exp\left(-\frac{\mu_V^2}{2\sigma_V^2}\right) + o(\theta_V), \quad (62)$$

Therefore, the asymptotic result for TPLR can be further written as

$$\text{TPLR} \simeq \frac{\exp\left(-\frac{\lambda}{2}\right)}{\Gamma(1.5)\sqrt{2}} \exp\left(-\frac{\mu_V^2}{2\sigma_V^2}\right) \frac{\int_{-\sqrt{\Theta_D}}^{\sqrt{\Theta_D}} \left(\frac{\Theta_D - \theta_V^2}{\sigma_H^2}\right)^{\frac{1}{2}} d\theta_V}{\sqrt{2\pi\sigma_V^2}}, \quad (63)$$

where the integral can be easily solved in a closed-form,

$$\int_{-\sqrt{\Theta_D}}^{\sqrt{\Theta_D}} \left(\frac{\Theta_D - \theta_V^2}{\sigma_H^2}\right)^{\frac{1}{2}} d\theta_V = \frac{\pi}{2\sigma_H} \Theta_D. \quad (64)$$

Finally, the asymptotic expression for TPLR is derived as (10).

APPENDIX D: PROOF OF THEOREM 2

The TPE can be written as

$$\text{TPE} = \Pr\left\{\theta_V^2 + \alpha\theta_V + \theta_H^2 \geq \Theta_E\right\} = \Pr\left\{\frac{\theta_H^2}{\sigma_H^2} \geq \frac{\Theta_E - \theta_V^2 - \alpha\theta_V}{\sigma_H^2}\right\}, \quad (65)$$

where θ_H^2/σ_H^2 follows the non-central Chi-squared distribution with the unit degree and non-centrality parameter λ . By using the CDF of θ_H^2/σ_H^2 , we can get

$$\begin{aligned} \text{TPE} &= \mathbb{E}_{\theta_V} \left\{ Q_{\frac{1}{2}} \left(\sqrt{\lambda}, \sqrt{\frac{\Theta_E - \theta_V^2 - \alpha\theta_V}{\sigma_H^2}} \mathbb{I}\{\Theta_E \geq \theta_V^2 + \alpha\theta_V\} \right) \right\} \\ &= \frac{1}{\sqrt{2\pi\sigma_V^2}} \int_{-\infty}^{+\infty} \exp\left(-\frac{(\theta_V - \mu_V)^2}{2\sigma_V^2}\right) Q_{\frac{1}{2}} \left(\sqrt{\lambda}, \sqrt{\frac{\Theta_E - \theta_V^2 - \alpha\theta_V}{\sigma_H^2}} \mathbb{I}\{\Theta_E \geq \theta_V^2 + \alpha\theta_V\} \right) d\theta_V. \end{aligned} \quad (66)$$

Let $x = \frac{\theta_V - \mu_V}{\sqrt{2\sigma_V}}$ in (66), and then we can get the following integral form,

$$\text{TPE} = \frac{1}{\sqrt{\pi}} \int_{-\infty}^{+\infty} \exp(-x^2) f_{\text{TPE}}(x) dx, \quad (67)$$

which can be easily approximated by using the GHQ method, shown in Theorem 2.

APPENDIX E: PROOF OF LEMMA 2

By using the asymptotic expression for $F_{\theta_H^2/\sigma_H^2}(\cdot)$, the TPE can be written as

$$\text{TPE} \simeq 1 - \frac{\exp\left(-\frac{\lambda}{2}\right)}{\sqrt{2}\Gamma(1.5)} \mathbb{E}_{\theta_V} \left\{ \left(\frac{\Theta_E - \theta_V^2 - \alpha\theta_V}{\sigma_H^2} \right)^{\frac{1}{2}} \mathbb{I}\{\Theta_E \geq \theta_V^2 + \alpha\theta_V\} \right\}, \quad (68)$$

where

$$\begin{aligned} &\mathbb{E}_{\theta_V} \left\{ \left(\frac{\Theta_E - \theta_V^2 - \alpha\theta_V}{\sigma_H^2} \right)^{\frac{1}{2}} \mathbb{I}\{\Theta_E \geq \theta_V^2 + \alpha\theta_V\} \right\} \\ &= \frac{1}{\sqrt{2\pi\sigma_V^2}} \int_{-\infty}^{+\infty} \exp\left(-\frac{(\theta_V - \mu_V)^2}{2\sigma_V^2}\right) \left(\frac{\Theta_E - \theta_V^2 - \alpha\theta_V}{\sigma_H^2} \right)^{\frac{1}{2}} \mathbb{I}\{\Theta_E \geq \theta_V^2 + \alpha\theta_V\} d\theta_V. \end{aligned} \quad (69)$$

To have a possible integral interval, the indicator function shows that we must have

$$\theta_V^2 + \alpha\theta_V - \Theta_E \leq 0. \quad (70)$$

To have real solutions for the above inequality, we have

$$\alpha^2 + 4\Theta_E \geq 0 \implies \Theta_E \geq \frac{-\alpha^2}{4}. \quad (71)$$

If $\Theta_E < -\frac{\alpha^2}{4}$, the indicator function in (69) is always equal to zero. Therefore, the TPE is always 1, which is straightforward. In the following, we consider the case of $\Theta_E \geq -\frac{\alpha^2}{4}$. Let

$$\theta_1 = \frac{-\alpha + \sqrt{\alpha^2 + 4\Theta_E}}{2}, \quad \theta_2 = \frac{-\alpha - \sqrt{\alpha^2 + 4\Theta_E}}{2}.$$

The integral in (69) can be derived as (72),

$$\begin{aligned} & \mathbb{I} \left\{ \Theta_E \geq \frac{-\alpha^2}{4} \right\} \int_{\theta_2}^{\theta_1} \exp \left(-\frac{(\theta_V - \mu_V)^2}{2\sigma_V^2} \right) \left(\frac{\Theta_E - \theta_V^2 - \alpha\theta_V}{\sigma_H^2} \right)^{\frac{1}{2}} d\theta_V \\ & \stackrel{\Theta_E \rightarrow -\frac{\alpha^2}{4}}{\simeq} \mathbb{I} \left\{ \Theta_E \geq \frac{-\alpha^2}{4} \right\} \exp \left(-\frac{(\alpha/2 + \mu_V)^2}{2\sigma_V^2} \right) \int_{\theta_2}^{\theta_1} \left(\frac{\Theta_E - \theta_V^2 - \alpha\theta_V}{\sigma_H^2} \right)^{\frac{1}{2}} d\theta_V \\ & = \mathbb{I} \left\{ \Theta_E \geq \frac{-\alpha^2}{4} \right\} \exp \left(-\frac{(\alpha/2 + \mu_V)^2}{2\sigma_V^2} \right) \frac{\pi(\alpha^2 + 4\Theta_E)}{8\sigma_H}. \end{aligned} \quad (72)$$

Finally, combining (68) and (72) yields (35).

APPENDIX F: PROOF OF THEOREM 3

Using the iterative expectation method, i.e., $\mathbb{E}\{XY\} = \mathbb{E}_Y\{\mathbb{E}_{X|Y}(X|Y)\}$ in (37), we have

$$\begin{aligned} \text{TPRE} &= \mathbb{E}_{\theta_V} \left\{ \mathbb{I} \left\{ \frac{\Theta_E - \theta_V^2 - \alpha\theta_V}{\sigma_H^2} \leq \frac{\Theta_D - \theta_V^2}{\sigma_H^2} \right\} \right. \\ & \quad \left. \times \left[F_{\theta_H^2/\sigma_H^2} \left(\frac{\Theta_D - \theta_V^2}{\sigma_H^2} \right) - F_{\theta_H^2/\sigma_H^2} \left(\frac{\Theta_E - \theta_V^2 - \alpha\theta_V}{\sigma_H^2} \right) \right] \right\}, \end{aligned} \quad (73)$$

where

$$\begin{aligned} & F_{\theta_H^2/\sigma_H^2} \left(\frac{\Theta_D - \theta_V^2}{\sigma_H^2} \right) - F_{\theta_H^2/\sigma_H^2} \left(\frac{\Theta_E - \theta_V^2 - \alpha\theta_V}{\sigma_H^2} \right) \\ & = Q_{\frac{1}{2}} \left(\sqrt{\lambda}, \sqrt{\frac{\Theta_E - \theta_V^2 - \alpha\theta_V}{\sigma_H^2} \mathbb{I} \{ \Theta_E \geq \theta_V^2 + \alpha\theta_V \}} \right) - Q_{\frac{1}{2}} \left(\sqrt{\lambda}, \sqrt{\frac{\Theta_D - \theta_V^2}{\sigma_H^2} \mathbb{I} \{ \Theta_D \geq \theta_V^2 \}} \right), \end{aligned} \quad (74)$$

and $\mathbb{I}\left\{\frac{\Theta_E - \theta_V^2 - \alpha\theta_V}{\sigma_H^2} \leq \frac{\Theta_D - \theta_V^2}{\sigma_H^2}\right\} = \mathbb{I}\left\{\theta_V \geq \frac{\Theta_E - \Theta_D}{\alpha}\right\}$. The TPRE is further written as (75),

$$\begin{aligned} \text{TPRE} = & \frac{1}{\sqrt{2\pi\sigma_V^2}} \int_{-\infty}^{+\infty} \left[Q_{\frac{1}{2}} \left(\sqrt{\lambda}, \sqrt{\left(\frac{\Theta_E - \theta_V^2 - \alpha\theta_V}{\sigma_H^2}\right)} \mathbb{I}\left\{\Theta_E \geq \theta_V^2 + \alpha\theta_V\right\} \right) \right. \\ & \left. - Q_{\frac{1}{2}} \left(\sqrt{\lambda}, \sqrt{\frac{\Theta_D - \theta_V^2}{\sigma_H^2}} \mathbb{I}\left\{\Theta_D \geq \theta_V^2\right\} \right) \right] \exp\left(-\frac{(\theta_V - \mu_V)^2}{2\sigma_V^2}\right) \mathbb{I}\left\{\theta_V \geq \frac{\Theta_E - \Theta_D}{\alpha}\right\} d\theta_V. \end{aligned} \quad (75)$$

Let $x = \frac{\theta_V - \mu_V}{\sqrt{2}\sigma_V}$, and then, the integral form becomes (76),

$$\begin{aligned} \text{TPRE} = & \frac{1}{\sqrt{\pi}} \int_{-\infty}^{+\infty} \exp(-x^2) \mathbb{I}\left\{x \geq \frac{\Theta_E - \Theta_D}{\alpha\sqrt{2}\sigma_V} - \frac{\mu_V}{\sqrt{2}\sigma_V}\right\} \\ & \times \left[Q_{\frac{1}{2}} \left(\sqrt{\lambda}, \frac{1}{\sigma_H} \sqrt{\Theta_E - \left(\sqrt{2}\sigma_V x + \mu_V\right)^2 - \alpha\left(\sqrt{2}\sigma_V x + \mu_V\right)} \right) \right. \\ & \quad \left. \times \mathbb{I}\left\{\Theta_E \geq \left(\sqrt{2}\sigma_V x + \mu_V\right)^2 + \alpha\left(\sqrt{2}\sigma_V x + \mu_V\right)\right\} \right) \\ & \left. - Q_{\frac{1}{2}} \left(\sqrt{\lambda}, \frac{1}{\sigma_H} \sqrt{\left(\Theta_D - \left(\sqrt{2}\sigma_V x + \mu_V\right)^2\right)} \mathbb{I}\left\{\Theta_D \geq \left(\sqrt{2}\sigma_V x + \mu_V\right)^2\right\} \right) \right] dx. \end{aligned} \quad (76)$$

By using the GHQ method, an approximate result is shown in Theorem 3.

APPENDIX G: DERIVATION OF THE JOINT PDF $f_{X,Y}(\cdot, \cdot)$

Let $Z = \theta_H^2$. We have $X = \theta_V^2 + Z$ and $Y = 2\theta_V^2 + 2Z + 2\alpha\theta_V$. The PDF of Z is

$$f_Z(z) = \frac{1}{\sqrt{2\pi\sigma^2}} z^{-\frac{1}{2}} \exp\left(-\frac{z}{2\sigma^2}\right), \quad z \geq 0. \quad (77)$$

The Jacobian matrix is

$$\begin{vmatrix} \frac{\partial X}{\partial \theta_V} & \frac{\partial X}{\partial Z} \\ \frac{\partial Y}{\partial \theta_V} & \frac{\partial Y}{\partial Z} \end{vmatrix} = \begin{vmatrix} 2\theta_V & 1 \\ 4\theta_V + 2\alpha & 2 \end{vmatrix} = 4\theta_V - 4\theta_V - 2\alpha = -2\alpha. \quad (78)$$

The joint PDF of X and Y is finally derived as

$$\begin{aligned} f_{X,Y}(x,y) &= f_{\theta_V}(\theta_V) f_Z(z) \left| \frac{\partial(X,Y)}{\partial(\theta_V,Z)} \right|^{-1} = \frac{1}{2\alpha} f_{\theta_V}\left(\frac{y-2x}{2\alpha}\right) f_Z\left(x - \left(\frac{y-2x}{2\alpha}\right)^2\right) \\ &= \frac{1}{4\pi\sigma^2\alpha} \exp\left(-\frac{x}{2\sigma^2}\right) \left[x - \left(\frac{y-2x}{2\alpha}\right)^2 \right]^{-\frac{1}{2}} \mathbb{I}\left\{-2\alpha\sqrt{x} + 2x \leq y \leq 2x + 2\alpha\sqrt{x}\right\} \mathbb{I}\{x \geq 0\}. \end{aligned} \quad (79)$$

APPENDIX H: DERIVATION OF THE DOUBLE INTEGRAL IN TPRE

In order to simplify the notations, we define the double integral in (45) as \mathcal{I}_1 . By using the integral identity

$$\int_{\Theta_E}^{2x+2\alpha\sqrt{x}} \left[x - \left(\frac{y-2x}{2\alpha} \right)^2 \right]^{-\frac{1}{2}} dy = \alpha\pi - 2\alpha \arcsin \left(\frac{\Theta_E - 2x}{2\alpha\sqrt{x}} \right), \quad (80)$$

\mathcal{I}_1 can be written as

$$\mathcal{I}_1 = \int_0^{\Theta_D} \exp \left(-\frac{x}{2\sigma^2} \right) \left[\alpha\pi - 2\alpha \arcsin \left(\frac{\max \{ \Theta_E, -2\alpha\sqrt{x} + 2x \} - 2x}{2\alpha\sqrt{x}} \right) \right] dx, \quad (81)$$

which can be easily calculated by using Gaussian quadrature method, $\mathcal{I}_1 \approx \frac{\Theta_D}{2} \sum_{i=0}^N \omega'_i f'_{\text{TPRE}} \left(\frac{\Theta_D}{2} x'_i + \frac{\Theta_D}{2} \right)$.

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