

Abstract

A proof of convergence of the standard EnKF generalized to non-Gaussian state space models is provided. A density-based deterministic approximation of the mean-field limiting EnKF (MFEnKF) is proposed, consisting of a PDE solver and a quadrature rule. Given a certain minimal order of convergence κ between the two, this extends to the deterministic filter approximation, which is therefore asymptotically superior to standard EnKF for $d < 2\kappa$. The fidelity of approximation of the true distribution is also established using an extension of total variation metric to random measures. This is limited by a Gaussian bias term arising from non-linearity/non-Gaussianity of the model, which arises in both deterministic and standard EnKF. Numerical results support and extend the theory.

1. Setting

Let $\mathcal{K} : \mathbb{R}^d \rightarrow \mathcal{P}(\mathbb{R}^d)$. Consider the Markov chain

$$\begin{aligned} u_{j+1} &\sim \mathcal{K}(u_j, \cdot), \quad j \in \mathbb{N}, \\ u_0 &\sim N(m_0, C_0). \end{aligned} \quad (1)$$

Data $Y_k = \{y_j\}_{j=1}^k$ defined as

$$y_j = Hu_j + \eta_j, \quad j \in \mathbb{Z}^+,$$

where $H : \mathbb{R}^d \rightarrow \mathbb{R}^m$ is linear and $\eta = \{\eta_j\}_{j \in \mathbb{Z}^+}$ is an i.i.d. sequence, independent of u_0 and the noise in \mathcal{K} , with $\eta_1 \sim N(0, \Gamma)$.

1.1 Filtering distribution

Distribution of $u_j | Y_j$ is *gold standard*. Likelihood function is

$$g(u, y) = \frac{\tilde{g}(u, y)}{\int \tilde{g}(u, y) du}, \quad \tilde{g}(u, y) = e^{-\frac{1}{2} |y - Hu|^2}, \quad (2)$$

with shorthand $g_j(u) = g(u, y_j)$. Define \mathcal{C}_j as follows. For $u \sim \hat{\rho}$, $u | y_j \sim \rho$, where

$$\rho = \mathcal{C}_j \hat{\rho} := \frac{\hat{\rho} g_j}{\int \hat{\rho} g_j}. \quad (3)$$

Recursion for filtering density $\rho_j = \mathcal{C}_j \mathcal{K} \rho_{j-1}$

$$\text{Prediction } \hat{\rho}_j = \mathcal{K} \rho_{j-1},$$

$$\text{Update } \rho_j = \mathcal{C}_j \hat{\rho}_j,$$

Optimal filter given by

$$\mathbb{E}(u_j | Y_j) = \operatorname{argmin}_{\hat{u}_j(Y_j)} \mathbb{E}[\hat{u}_j(Y_j) - u_j]^2.$$

1.2 Optimal Linear Filtering

$$m_j(y_j) = \operatorname{argmin}_{\{\hat{u}_j(y_j) = K_j y_j + b_j\}} \mathbb{E}[\hat{u}_j(y_j) - u_j]^2,$$

Optimizing with respect to K_j and b_j gives

$$m_j(y_j) = \mathbb{E}u_j + K_j(y_j - H\mathbb{E}u_j),$$

$$K_j = \mathbb{E}[(u_j - \mathbb{E}u_j) \otimes (y_j - \mathbb{E}y_j)] \mathbb{E}[(y_j - \mathbb{E}y_j) \otimes (y_j - \mathbb{E}y_j)]^{-1}.$$

2. ENKF

2.1 Mean-field EnKF

$$\text{Prediction } \begin{cases} \hat{v}_{j+1} \sim \mathcal{K}(v_j, \cdot), \\ j+1 = \mathbb{E}\hat{v}_{j+1}, \\ \hat{C}_{j+1} = \mathbb{E}(\hat{v}_{j+1} - j+1) \otimes (\hat{v}_{j+1} - j+1) \end{cases}$$

$$\text{Analysis } \begin{cases} S_{j+1} = H\hat{C}_{j+1}H^T + \Gamma \\ K_{j+1} = \hat{C}_{j+1}H^T S_{j+1}^{-1} \\ \hat{y}_{j+1} = y_{j+1} + \eta_{j+1} \\ v_{j+1} = (I - K_{j+1}H)\hat{v}_{j+1} + K_{j+1}\hat{y}_{j+1}. \end{cases}$$

Here η_j are i.i.d. draws from $N(0, \Gamma)$

2.2 Finite ensemble EnKF

$$\text{Prediction } \begin{cases} \hat{v}_{j+1}^{(n)} \sim \mathcal{K}(v_j^{(n)}, \cdot), \quad n = 1, \dots, N \\ \hat{m}_{j+1} = \frac{1}{N} \sum_{n=1}^N \hat{v}_{j+1}^{(n)}, \\ \hat{C}_{j+1} = \frac{1}{N} \sum_{n=1}^N (\hat{v}_{j+1}^{(n)} - \hat{m}_{j+1}) \otimes (\hat{v}_{j+1}^{(n)} - \hat{m}_{j+1}) \end{cases}$$

$$\text{Analysis } \begin{cases} v_{j+1}^{(n)} = (I - K_{j+1}H)\hat{v}_{j+1}^{(n)} + K_{j+1}y_{j+1}^{(n)} \\ y_{j+1}^{(n)} = y_{j+1} + \eta_{j+1}^{(n)} \end{cases}$$

3. Fokker-Planck Filters

$$du = F(u)dt + \sqrt{2bd}W, \quad u(0) = u_0,$$

$F : \mathbb{R}^d \rightarrow \mathbb{R}^d$ differentiable and Lipschitz, $b \in (0, \infty)$ constant.

Density governed by Fokker-Planck equation

$$\partial_t \rho = \mathcal{G} \rho, \quad \mathcal{G} \rho = \nabla \cdot (b \nabla \rho - F \rho), \quad \rho(u, 0) = \delta(u_0 - u),$$

3.1 FPF Algorithms

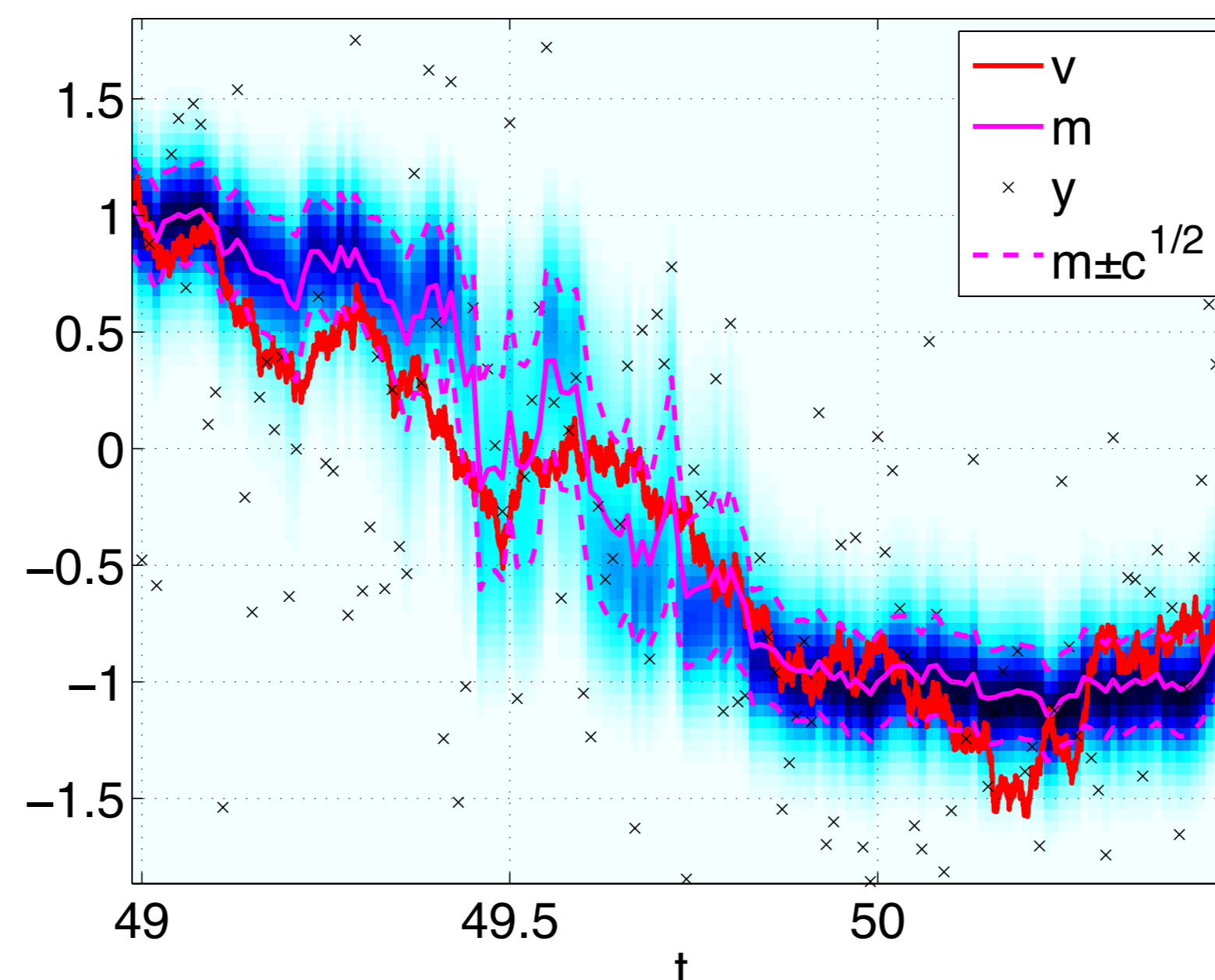


Figure 1: $F(u) = au(1-u^2)/(1+u^2)$, $m = \mathbb{E}(u|y)$, $c = \mathbb{E}[(u-m)^2|y]$.

Full FPF

- (1) Discretize the density at time j over space.
- (2) Evolve with accurate time-stepper to time $j+1$.
- (3) Update using integration rule, return to step (2).

MFEnKF

- (1), (2) Same as above.
- (3) Obtain mean and covariance of the prediction.
- (4) Change variables $u \rightarrow (I - K_j H)u$ using interpolation.
- (5) Convolve with the density of $K_j N(y_j, \Gamma)$ on range of K_j .

MFEnKF-G1

- (1),(2),(3) Same as in Algorithm Full FPF.
- (4) Approximate the updated distribution by the Gaussian.

MFEnKF-G2

- (1),(2) Same as above.
- (3) Approximate the predicting density by a Gaussian.
- (4) Use Kalman update formula for the mean and covariance.

Approximated in d -dimensions with N d.o.f. to $\mathcal{O}(N^{-\kappa/d})$: $\kappa = \min\{\kappa_1, \kappa_2\}$, PDE method order κ_1 , quadrature order κ_2 .

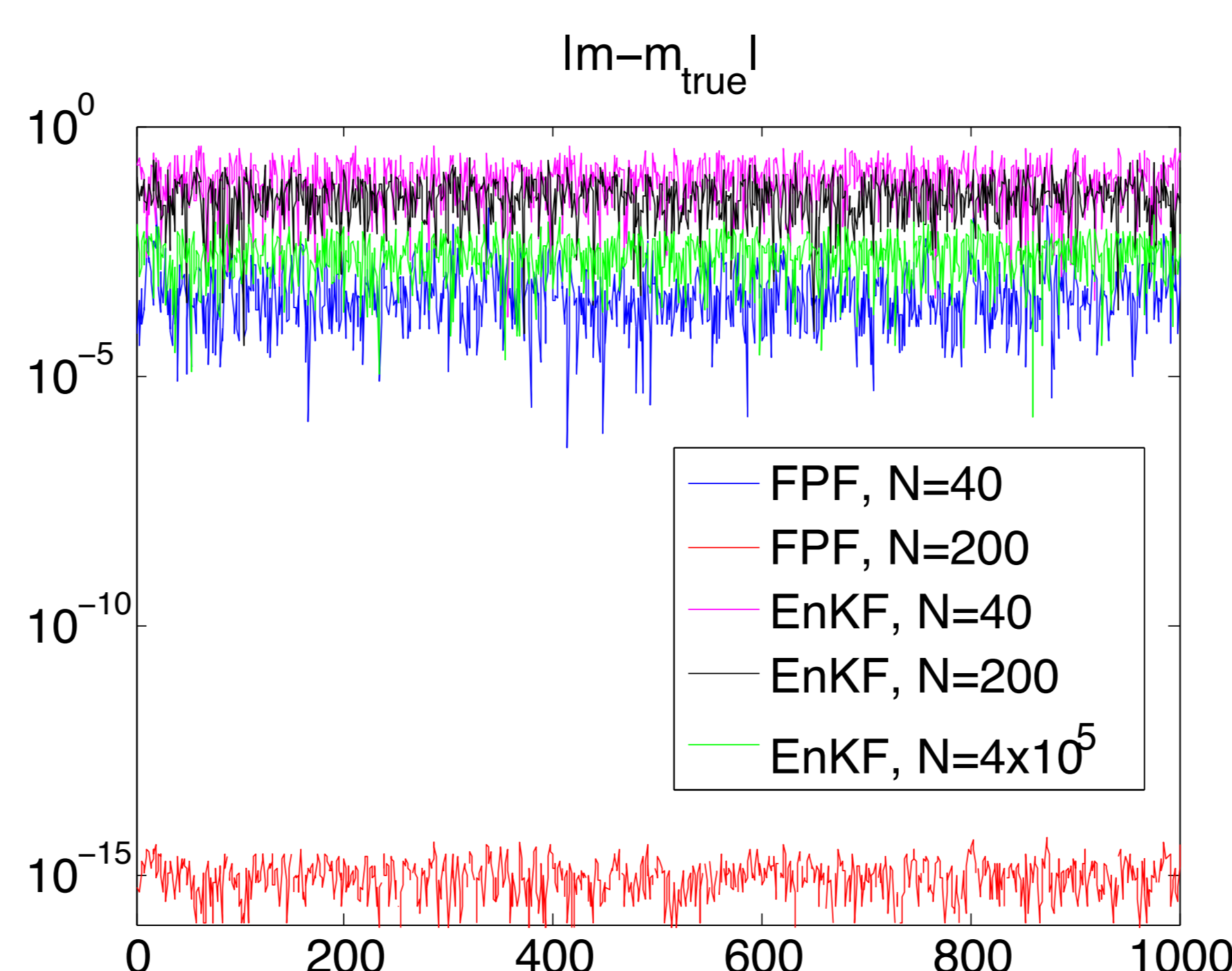


Figure 2: $F(u) = -au$, Error of FPF MFEnKF-G2 vs. EnKF.

3.2 ENKF Converges to MFENKF

Let $\pi, f : \mathbb{R}^d \rightarrow \mathbb{R}$ and $|f|_L = |f|_\infty + \sup_{x \neq x'} \frac{|f(x) - f(x')|}{|x - x'|}$. Define

$$\|\pi\| = \sup_{|f|_L \leq 1} \sqrt{\mathbb{E} \left| \int f \pi \right|^2},$$

Now define

$$d(\pi, \rho) = \|\pi - \rho\|. \quad (4)$$

- ρ_j – filtering density.
- π_j – the mean-field EnKF density.
- $\hat{\pi}_j^N \approx \pi_j$ – the deterministic approximation (MFEnKF).
- $\hat{\pi}_j^N \approx \pi_j$ – the standard Monte-Carlo EnKF.

THEOREM I: It is true that $d(\hat{\pi}_j^N, \pi_j) = \mathcal{O}(N^{-1/2})$.

Proof: See [1], Sec. 4.3.1, for extension of [2].

4. MFENKF is not the Posterior

$$\begin{aligned} d(\hat{\pi}_j^N, \rho_j) &\leq \underbrace{d(\hat{\pi}_j^N, \pi_j)}_{\text{ensemble error}} + \underbrace{d(\pi_j, \rho_j)}_{\text{Gaussian error}}, \\ d(\pi_j^N, \rho_j) &\leq \underbrace{d(\pi_j^N, \pi_j)}_{\text{discretization error}} + \underbrace{d(\pi_j, \rho_j)}_{\text{Gaussian error}}. \end{aligned}$$

Assume (i) $\alpha \leq g(u, y) \leq \alpha^{-1} \forall u \in \mathbb{R}^d, y \in \mathbb{R}^m, j \in \mathbb{Z}$. And (ii) for time h , $\mathcal{K}_h \pi_j = \pi_j + \mathcal{O}(h)$ and $\mathcal{K}_h \rho_j = \rho_j + \mathcal{O}(h)$. Then

THEOREM II: Given observation increment h , and under assumptions above, as $h \rightarrow 0$ and $N \rightarrow \infty$

$$d(\pi_j^N, \rho_j) = \mathcal{O}(N^{-\kappa/d} + \lambda h) \quad \text{and} \quad d(\hat{\pi}_j^N, \rho_j) = \mathcal{O}(N^{-1/2} + \lambda h)$$

where $\lambda = 0$ if F is linear and $\lambda = 1$ if f is nonlinear.

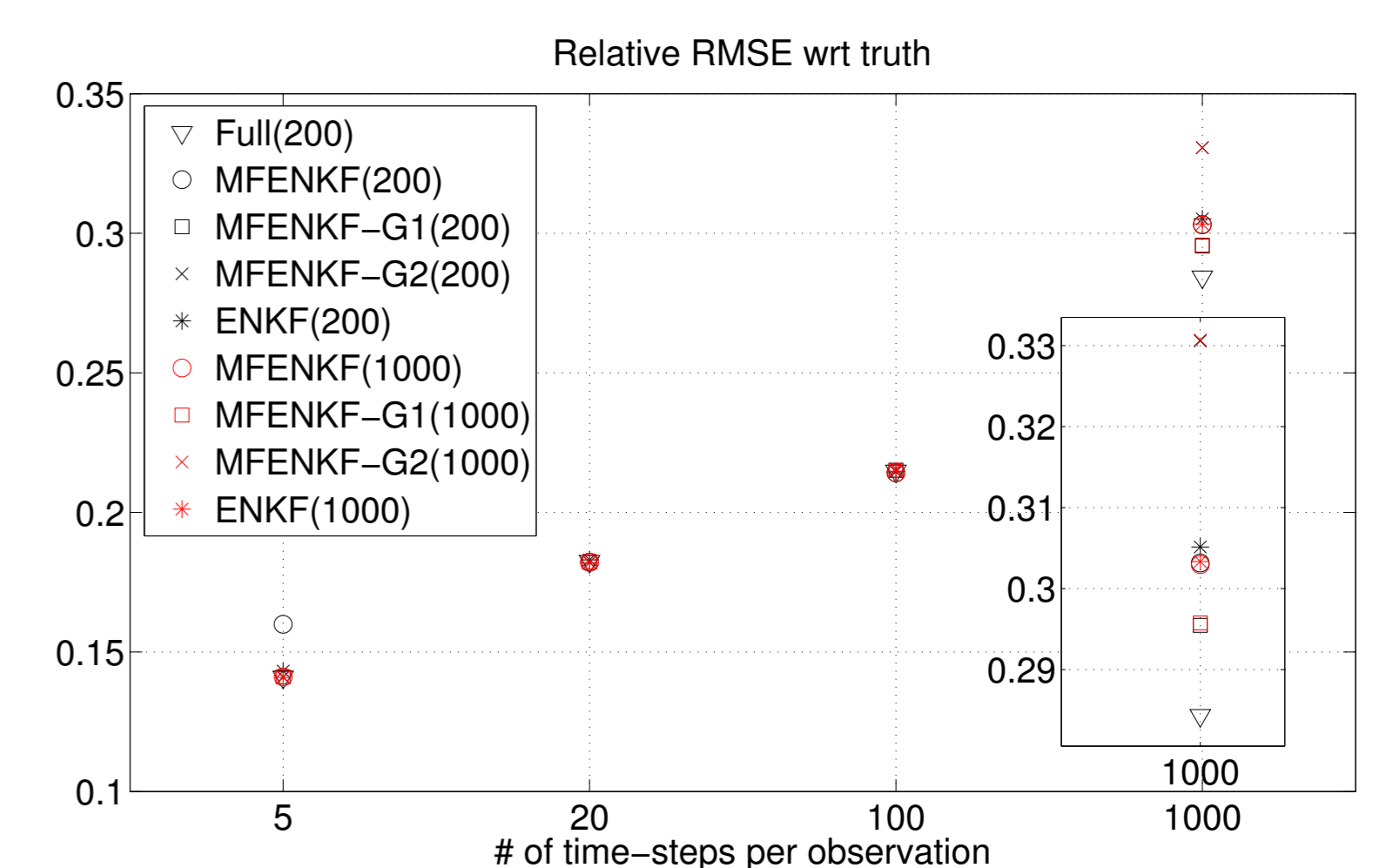


Figure 3: Nonlinear case, RMSE of mean with respect to the true (unconditioned) signal.

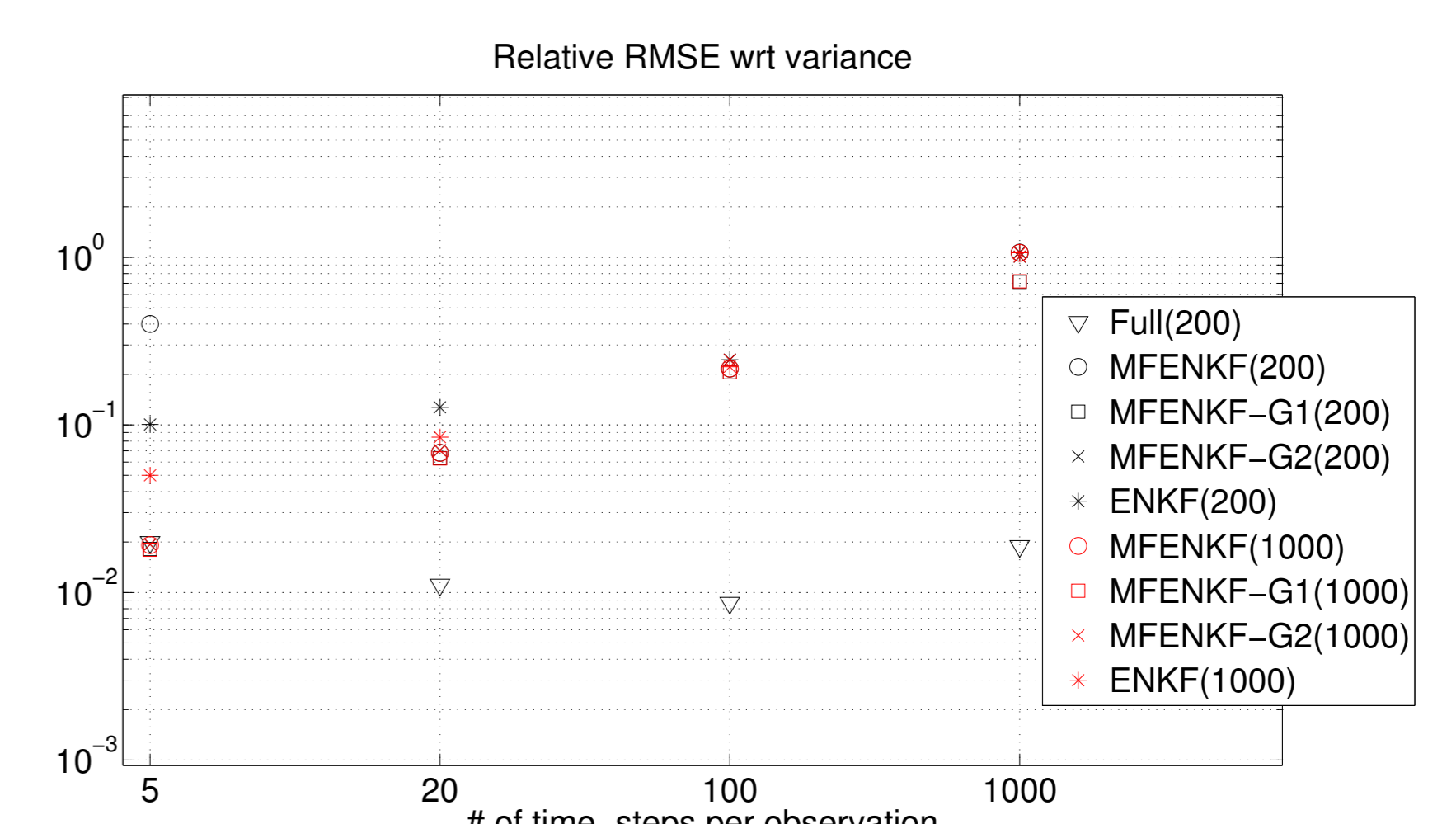


Figure 4: Nonlinear case, RMSE of covariance with respect to the true posterior covariance.

Acknowledgements

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References

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- [2] F. LeGland, V. Monbet, V.D. Tran. "Large sample asymptotics for the ensemble kalman filter." The Oxford Handbook of Nonlinear Filtering, 598-631, 2011.